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RELATION OF SICKNESS TO INCOME AND INCOME CHANGE IN 10 SURVEYED COMMUNITIES *

Health and Depression Studies No. 1: Method of Study and General Results for Each Locality

By G. St. J. Perrott, Consultant, and Selwin D. Collins, Senior Statistician United States Public Health Service

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The ordinary barometers of health—death rates and reports of communicable diseases—do not indicate that harmful effects of the depression upon the health of the population as a whole have taken place. The comfortable conclusion is drawn by many that the physical well-being of the American people not only has not suffered but, in view of the continued low death rate, may have been benefited

From the Office of Statistical Investigations, U. S. Public Health Service, and the Division of Research,
 Milbank Memorial Fund.

This study was made also in cooperation with the international inquiry being carried out in various countries under the general auspices of the health organization of the League of Nations, the members of the American committee being Edgar Sydenstricker, Milbank Memorial Fund; Louis I. Dublin, Metropolitan Life Insurance Co.; Walter F. Willcox, Cornell University; and Selwyn D. Collins, U. S. Public Health Service

This is the first of a series of papers on sickness and medical care among groups of white wage-earning families severely affected by unemployment during the economic depression. Preliminary papers, giving results for parts of the surveyed group, have been published as follows: Perrott, G. St. J., Collins, Selwyn D., and Sydenstricker, Edgar. Sickness and the economic depression, Public Health Reports, Oct. 13, 1933 (Reprint No. 1598). Perrott, G. St. J., and Collins, Selwyn D.: Sickness and the depression, Milbank Memorial Fund Quarterly Bulletin, October 1933, vol. 11, no. 4, pp. 261-298; January 1934, vol. 12, no. 1, pp. 28-34; July 1934, vol. 12, no. 3, pp. 218-224; American Journal of Public Health, February 1934, vol. 24, no. 2, pp. 101-107. Collins, Selwyn D., and Perrott, G. St. J.: The economic depression and sickness, Journal of the American Statistical Association, March 1934, Supplement 29, pp. 47-51. Perrott, G. St. J., Sydenstricker, Edgar, and Collins, Selwyn D.: Medical care during the depression, Milbank Memorial Fund Quarterly Bulletin, April 1934, vol. 12, no. 2, pp. 99-114. Sydenstricker, Edgar, and Perrott, G. St. J.: How unemployment affects illness and hospital care, The Modern Hospital, March 1934, vol. 42, no. 3, pp. 41-44.

¹ The death rate from all causes reached the lowest figure on record in the first half of 1933, but during the winter of 1933-34 mortality was on a slightly higher level than in corresponding months of immediately preceding years, except for periods in those years when influenza was epidemic. While the rise was slight, it is consistently evident in a large proportion of the 28 States for which preliminary figures are available. (See Public Health Reports, Nov. 9, 1934, Mortality from certain causes during the first half of 1931.)

by the economic catastrophe. Such a conclusion, based upon mortality statistics alone, is open to question. Even in the worst depression the families of the unemployed are a minority, and the trend of mortality in the total population does not necessarily reflect the trend in these severely affected households.

The assumption that mortality in the general population is an accurate index of sickness in the families of the unemployed is still less tenable. Recent morbidity studies 2 have shown that the important causes of death are not the most frequent causes of illness. The number of illnesses severe enough to be remembered and reported, even in relatively infrequent canvasses of households, is 75 to 100 times the number of deaths. For digestive, respiratory, eye, ear, and skin affections and the common communicable diseases of childhood, the disparity between sicknesses and deaths is even greater. In depending upon deaths to indicate trends in health we are relying on a small and probably biased sample of the cases of illness. The desirability of checking up on all illnesses before drawing conclusions from data based only on the fatal cases seems apparent.

Among the now well-recognized indexes of ill health are records of sickness. When properly obtained and analyzed, they reveal some of the reactions of human beings to immediate environmental factors in a far more sensitive degree than the gross death rate or even mortality by cause can possibly do. Since no national system for the complete registration of sickness exists, special records must be collected, a difficulty not without its advantages, since it permits information to be obtained for such groups and in such detail as may be desired. One phase of the study of health and the depression by the Public Health Service and the Milbank Memorial Fund utilized this method extensively. A sickness and mortality survey was made in 1933 of nearly 12,000 wage-earning families which had suffered from the depression in varying degrees of severity. Among the more specific purposes of the study were the following:

1. To ascertain whether or not there is any association between income changes during the depression and ill health as measured by morbidity and mortality.

2. If such an association exists, to discover what kinds of sickness

are chiefly responsible for the association.

3. To determine the amount and kinds of medical care received by various economic groups of the people.

4. To study diets and housing conditions of selected families among

the employed and the unemployed.

5. Using school records of height and weight, to study the growth of children in families of the "new poor" in the surveyed households as compared with children in families that remained in comfortable circumstances throughout the depression.

² Hagerstown Morbidity Studies, the Public Health Reports for Feb. 13, 1925, and June 14, 1927 (Reprints 989 and 1167), respectively; Morbidity in 18 States, Public Health Reports for Mar. 24, 1933 (reprint 1563), and Publication No. 27 of the Committee on the Costs of Medical Care, University of Chicago Press, 1983.

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METHOD AND SCOPE OF SURVEY

These included eight large cities—Baltimore, Birmingham, Brooklyn, Cleveland, Detroit, New York (Borough of Manhattan), Pittsburgh, and Syracuse, a group of coal mining communities in the vicinity of Morgantown, W. Va., and a group of cotton-mill villages in the vicinity of Greenville, S. C. About 1,200 families were visited in each locality.

No attempt was made to select sections that would be representative of the city as a whole; only the poorer districts were canvassed. Slum areas were not included, because they would contain too many families who had never, even at the height of prosperity, been self-supporting. Well-to-do sections were omitted as being still above a standard of living that could affect health adversely, even though great decreases in income had taken place. Colored sections were excluded to avoid the question of racial differences in employment, income, and sickness. In blocks or streets that were surveyed, every white family was included, whether employed or unemployed and whether recently poor or never self-supporting. Those families whose breadwinners still had their jobs were to serve an important role in the study, viz, as a control group whose illness rate would be a yardstick which would be essential in interpreting the illness rates found for those who had suffered economic reverses.

Previous experience in sickness surveys indicates that a single interview of a housewife will not yield a reasonably complete record of illness for a longer period than about 3 months. Even for that period, one cannot expect to get all of the many minor respiratory and digestive conditions that caused no disability but would be reported as illness if visits were made at weekly or semimonthly intervals. With this limitation on the illness record that could be secured, the problem was to plan a survey, with only one visit to the households, that would nevertheless afford more than a comparison of illness rates among poor and comfortable or among employed and unemployed at or immediately preceding the time of the canvass. A feasible method seemed to be to obtain for each member of the family (1) a record of illness and medical care for the 3 months preceding the date of the canvass, and (2) a record of occupation, wages earned, and regularity of employment for each year from 1929 to 1932 of sufficient detail to compute the family income. These data enable us to relate current illness to changes in income during the depression as well as to present economic and employment status. The accuracy of the 4-year income record may be doubted; but this was a period of such tremendous changes in economic well-being that small errors did not interfere with a reasonably good classification of the families according to income change since 1929.

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Although the enumerators were hired locally, the canvass in each city was in immediate charge of a person trained in the collection and tabulation of such data, who was assigned from the permanent personnel of the Public Health Service or the Milbank Memorial Fund. Because of the prevailing economic conditions it was possible to get exceptionally good enumerators. These enumerators canvassed families only after they had received careful instruction and had made trial visits with the local supervisor. All persons worked under uniform written instructions. Thoroughness, rather than speed, was encouraged in the enumerators. One of us (G. S. P.) acted as general supervisor and visited all but two of the communities either to start the work (select districts, enumerators, etc.) or to check the selections made by the local supervisor.

THE POPULATION SURVEYED

Number.—In the 10 localities, schedules were obtained from about 12,000 families. The data from 11,511 of these families, including 49,136 individuals, were finally coded and transferred to punch cards, and the remainder were discarded because of incompleteness of information on the schedule. In table 1 the percentage distribution of families in each locality is given according to nativity, occupation, employment status, and relief status. Only those families are included on which economic data were complete for the 4 years, 1929–32, as the major part of the sickness tabulations refer to this group.²

Nativity.—Considering the 8 large cities, in 40 percent of the families the male household head was native white of native parents, in 18 percent of foreign or mixed parents, and in 42 percent foreign born.³ The nativity of family heads varied considerably from city to city. Birmingham and Greenville were largely native white of native parents (95 and 100 percent, respectively), while in New York and Cleveland 60 percent of the family heads were foreign born (18 and 22 percent, respectively, native white of native parents). The racial stock of the group of foreign or mixed parents was largely English, Irish, and German, while that of the foreign-born group was more evenly distributed between English, Irish, Italian, Polish, and Slavic.

² Incomplete economic data prevented the use of 1,657 families in tabulations in which income classifications were made; 727 families whose heads were married since 1929 were omitted from tabulations where families were grouped by change in income between 1929 and 1932. This left a total of 9,127 families, including 40,184 individuals, in the 10 surveyed localities, on which economic data were complete for the 4 years and other information was reasonably detailed also. These families were used in all tabulations for the localities considered separately, when classification was made by income. For many tabulations the 8 large cities were combined into one group which comprised 7,436 families, including 31,635 individuals. The entire group of 11,511 families has been used in showing the association between illness and unemployment in 1932.

³ While no attempt was made to secure sample populations representative of the city, the nativity of the heads of surveyed families is similar to that of the 1930 census for each city (excluding Negroes) with the exception of Brooklyn and Syracuse. If the census data for each city are weighted by the number of families in the surveyed population, the average so obtained gives 40 percent native white of native parents, 23-percent native white of foreign or mixed parents and 37 percent foreign born, as compared with percent ages of 40, 18, and 42, respectively (see table 1), which were actually found in the surveyed families.

Table 1.—Percentage distribution of white wage-earning families 1 by (1) nativity of household head, (2) occupational status of chief wage earners in family in 1929 and 1932, and (4) families on relief at any time during 1932

			Total number	lies ob- served 1		961 780 780 1, 047 1, 008 1, 225 780 895 895 895 739 739 739 7436
				Fami- lies on relief		41.48851860
					full- time	48588888888 44
			22	One or	nore part- time, no full- time	8888888888
,		ly 3	1932	poloyed	Other families with no workers	œ 5 7 7 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9
	sification	Wage earners in family		All unemployed	Family Other has in- families come or with no pension workers	88 × 1 × 8 × 5 − 1 − 1 × 5 €
	filed class	e earner			full- time	888824288888
	Percentage distribution of families according to specified classification	Wag	æ	One or	more part- time, no full- time	00782222842 44
	accordin		1929	All unemployed	Family Other has in-families come or with no pension workers	20.0 0.6
	families				Family Other has in-families come or with no pension workers	#####################################
	ution of	chief			rercent unem- ployed, 1932	88 88 88 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8
	e distrib	Occupational status of chief	arner 2		Un- skilled labor	26 28 28 28 28 28 28 28 28 28 28 28 28 28
,	ercentag	oational	wage earner ?	Usual or 1929 occupation	Skilled labor	88888888888888888888888888888888888888
	F	Occul		Usual or	White- collar	38 38 38 10 10 10 10 10 10 10 10 10 10 10 10 10
		seho1d			For- eign- born	52 28 28 28 28 28 28 28 28 28 28 28 28 28
		rity of hou	head		Native— foreign parents	41 33 33 33 35 52 52 52 52 52 52 52 52 53 54 54 55 54 54 54 54 54 54 54 54 54 54
		Nativ			Native— native parents	25 25 25 25 25 25 25 25 25 25 25 25 25 2
		-	Loonlite	Sales of the sales		Baltimore Birmingham Brooklyn Cleveland Detroit New York Pittsburgh Syracuse Greenville Morgantown Total, 10 localities 4

unit under observation in 19.2.

Built under observation in 19.2.

Excludes unknown occupations. The term "white-collar" is here used to include all workers other than skilled and unskilled laborers, that is, professional, proprietary, and security and security includes "semiskilled". Farm laborers were present to a negligible extent and have been included with unskilled laborers. Household heads living on income of position are not included with the unemployed in 1932 and are excluded from the population in making this computation.

* Welgare work, when the sole occupation, was considered "unemployed."

* Weighted average. Excludes Greenville and Morgantown. ¹ Excludes 1,657 families for which economic data were incomplete and 727 families where marriage took place in 1930 or later. These are excluded also from tables 2, 3, 5, 6, and 7, but are included in table 4. The newly married families could not be used in tablations dealing with illness and income change, 1929-32, because they were not an economic

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Occupation.—The population was largely of the wage-earning class. In the 8 large cities the usual occupation of the chief wage earner was that of skilled or semiskilled laborer in 58.1 percent of the families; unskilled, 20.5 percent; clerical and kindred worker, 12.0 percent; proprietor, manager, or official, 7.8 percent; professional, 1.6 percent. In 1932 in 17 percent of the families the chief wage earner was without employment throughout the year. This figure varied from 6 percent in Brooklyn to 28 percent in Cleveland. In Greenville and Morgantown only 1 to 2 percent of the chief wage earners were unemployed in 1932. This low figure was due to the fact that only families having workers employed in the mills or mines were allowed to live in these company-owned villages.

Table 2.—Occupation shifts of chief wage earners between 1929 and 1932 in white families in 8 large cities

	Num-	Perce	entage of		ge earne oup in 19		occupa	tional
Occupation of household head in 1929	ber of fami- lies !	Unem- ployed	Pro- fes- sional	Pro- prie- tary	Cleri- cal	Skilled	Un- skilled	Total, all occu- tions, 1932
Professional Proprietary Clerical Skilled Unskilled	109 532 814 3, 946 1, 389	5. 5 8. 6 6. 1 17. 8 24. 9	.1 .1 .1	1.8 83.9 1.1 .9	0.9 3.4 87.1 .6	3. 0 3. 0 76. 4 1. 9	0.9 2.1 2.6 4.2 71.7	100. 0 100. 0 100. 0 100. 0 100. 0
All occupations	6, 790	16. 9	1. 5	7. 3	11. 2	45. 4	17. 7	100. 0

¹ Excludes families in which chief wage earner lived on income or pension in 1929 or 1932, families in which chief wage earner died after 1929, and families in which occupation of chief wage earner in 1929 or 1932 was unknown.

Unemployment and the shift in occupations between 1929 and 1932 are shown in table 2. Unemployment was highest among the unskilled laborers (25 percent) and lowest among the professional class (5.5 percent). Among skilled and unskilled laborers, the greatest shift was into the unemployed group, while in the clerical and proprietary classes, those who changed occupational status between 1929 and 1932 were about equally divided between the group that became unemployed and the groups that found other occupations. For example, 72 percent of the unskilled laborers were employed in the same class of occupation in 1932, 25 percent were unemployed, and 3 percent were in different occupational groups; 83 percent of the proprietary

⁴ Gainful white workers in the United States in 1930 similarly classified (excluding farm owners, tenants, and laborers) are distributed approximately as follows: Skilled and semiskilled, 39 percent; unskilled, 20 percent; clerks and kindred workers, 23 percent; proprietors, managers, and officials, 10 percent; professional workers, 8 percent. While the figures are not strictly comparable since the data of the present survey give the distribution of families by occupation of the chief wage earner, they indicate that the surveyed population contains an excess of skilled laborers and a deficiency of clerks and professional workers, as compared with the general population of the United States. See Edwards, Alba M.: A Social-Economic Grouping of the Gainful Workers in the United States. Journal American Statistical Association, December 1933, vol. 28, pp. 377-387.

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class remained in that category in 1932, 9 percent were unemployed, and 8 percent were in the clerical, skilled, and unskilled classes.

Employment status.—Considering all wage earners in the family, the data (table 1) show that in 1929 only 0.8 percent of the families in the 8 large cities had no employed workers, 14 percent had one or more part-time workers and no full-time workers, 82 percent had one or more full-time workers, with or without part-time workers, and 3 percent had wage earners living on income or pension. In 1932 there were 10 percent with no employed workers, 36 percent with part-time workers only, 48 percent with full-time workers, and 6 percent with wage earners living on income or pension. In 1932, 20 percent of all surveyed families were on public or private relief for part or all of the year. This proportion varied from 4 percent in Brooklyn to 30 percent in Pittsburgh.

Greenville and Morgantown presented an entirely different picture, with 72 percent of the families having part-time workers only, 28 percent having full-time workers, and no families having all workers unemployed. The reasons for this different showing have been discussed in a preceding paragraph.

Economic history of families.—Income as computed in this study includes all receipts from any source—wages, rents, interest, and profits, and also the amount of savings or borrowed funds used and the value of a food ticket or other receipts from public or private relief agencies. The figures for 1929, when only 4 percent of the families used savings or borrowed funds, represent income in the accepted sense of the word and may exceed expenditures; the figures for 1932, when about 20 percent of the families augmented their purchasing power by some use of savings or borrowed funds, are more properly called expenditures. This definition of income was adopted because it was desired to relate incidence of illness to standard of living, as expressed by expenditures rather than by actual income.

No attempt was made to select districts in which the income distribution of the surveyed families would be representative of the city as a whole. The plan, as already outlined, was to include sections having families that, in normal times, were in moderate circumstances, but that in large numbers had been reduced to poverty during the depression.

In table 3 the distribution of families in the 8 large cities by total income is shown for each year from 1929 to 1932, and for comparison the income as estimated for all nonfarm families in the United States.

The mean income of the surveyed group in 1929 was \$1,830, as compared with \$3,225 for the United States. The median income, which affords a better comparison, was \$1,650 in the surveyed group and \$1,900 for nonfarm families in the United States. If families with incomes above \$4,000 are excluded (these constitute 15 percent

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of the nonfarm families in the country), the income distribution of the surveyed group in 1929 is not far different from that of the nonfarm in the United States.⁵ By 1932, the median income of the surveyed group was \$870, which is a drop of 47 percent. In 1929, 26 percent of the canvassed families had incomes less than \$1,200 per year, as compared with 66 percent in 1932. On the other side of the picture, 35 percent of the families had incomes over \$2,000 in 1929 as compared with 10 percent in 1932.

Table 3.—Percentage distribution according to total income of families (1) in the surveyed population in 8 cities for 1929, 1930, 1931, and 1932, and (2) as estimated for the United States in 1929

Total family income per year	8	durveyed gro	up in 8 cities	1	Nonfarm families United States ²
•	1929	1930	1931	1932	1929
Under \$600.	6. 9	12. 4	20. 9	32. 4	4.0
\$600, but under \$1,200.	19. 5	25. 5	31. 0	33. 7	17.4
\$1,200, but under \$2,000.	38. 5	35. 2	30. 0	23. 4	32.0
\$2,000, but under \$3,000.	24. 2	19. 0	13. 5	8. 0	21.1
\$3,000, but under \$4,000.	7. 3	5. 4	3. 2	1. 7	10.2
\$4,000 and over.	3. 6	2. 5	1. 4	. 8	15.3
Total, all incomes.	100. 0	100. 0	100. 0	100. 0	100.0
Number of families Median income Mean income	7, 436	7, 436	7, 436	7, 436	21, 674, 000
	\$1, 650	\$1, 440	\$1, 160	\$870	\$1, 900
	1, 830	1, 600	1, 325	1, 050	3, 225

Baltimore, Birmingham, Brooklyn, Cleveland, Detroit, New York, Pittsburgh, and Syracuse.
 America's Capacity to Consume. By Maurice Leven, Harold G. Moulton, and Clark Warburton.
 The Brookings Institution, Washington, D. C., 1934.

The change from one income class to another is better shown in table 4, which indicates the correlation between 1929 and 1932 income. For example, in the group of families having less than \$600 annual income in 1929, 80 percent were still in that class in 1932. In the group having incomes between \$2,000 and \$3,000 in 1929, 17.5 percent were still in that class in 1932, 1 percent had risen to higher brackets, and the remainder had fallen into lower income groups.

The table suggests a means of classifying families according to economic experience, which is used later in relating sickness to *change* in income during the depression. For example, the group of families with less than \$600 annual income in 1932 constituted 32 percent of the surveyed group in the 8 large cities. Of this group, only 17 percent had been in this class in 1929, 66 percent had incomes between \$600 and \$2,000, and 17 percent had incomes over \$2,000 in 1929. In this study of illness as related to income change, we are particularly interested in 3 general classes of the population: (1) Families re-

⁴ The relatively high mean income (\$3,225) in the nonfarm families in the United States is due mainly to the families in the group above \$4,000, which constitute 15 percent of the families but receive 50 percent of the total income. In contrast, families receiving incomes over \$4,000 are less than 4 percent of the surveyed group and receive about 10 percent of the total income. This is reflected in the fact that while the mean income of nonfarm families in the United States was 75 percent higher, the median income was only 12 percent higher than that of the surveyed group in 1929.

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Table 4.—Income distribution in 1932 of families in 8 ¹ cities classified in 6 groups according to 1929 income

		Percen		milies in in the s				29 which
Annual family income in 1929	Num- ber of families	Under \$600	\$600 but under \$1,200	\$1,200 but under \$2,000	\$2,000 but under \$3,000	\$3,000 but under \$4,000	\$4,000 and over	Total, all incomes 1932
Under \$600	514	80. 2	17. 5	1.9	0.4			100.0
\$690 but under \$1,200	1, 450	49. 6	43, 6	6.3	. 5			100.0
\$1,200 but under \$2,000	2,860	31.0	39. 3	27.6	1.9	0. 2		100.0
\$2,000 but under \$3,000	1,801	17. 5	29. 3	34.6	17.5	1.0	0.1	100.0
\$3,000 but under \$4,000	540	10. 7	18. 9	29.4	28. 2	11, 5	1.3	100.0
\$4,000 and over	271	6. 6	12.5	24. 0	24. 4	14. 4	18, 1	100. 0
All incomes 1929	7, 436	32. 4	33, 7	23, 4	8.0	1.7	.8	100.0

¹ Baltimore, Birmingham, Brooklyn, Cleveland, Detroit, New York, Pittsburgh, and Syracuse.

maining in reasonably comfortable circumstances throughout the 4 years; (2) families that suffered material loss of income and, hence, lowered standard of living during the depression; and (3) families that were poverty-stricken even in 1929—the chronic poor. The first and third groups serve as controls, whose illness rates are compared with those of families that had suffered economic reverses.

DEFINITION OF ILLNESS AND METHOD OF CLASSIFYING

Inquiry was made about illness from all diseases and accidents, including mild as well as severe cases. What was included as illness was, to a considerable extent, a matter of what the informant (usually the housewife) remembered and designated as such. Hence the records of disabling cases are probably a better measure of real sickness than are the total cases, because the disabling illnesses are more likely to be accurately and completely reported. A case sufficiently severe to be disabling or confine the individual to his bed within 3 months of the interview is very likely to be remembered, while many of the minor ailments are forgotten and are consequently not mentioned to the enumerator.

The illness rates are for the 3-month period of the survey and are not reduced to an annual basis. All rates are adjusted for differences in age distribution. The "survey period" refers to the 3 months prior to the enumerator's visit; it is the period of time for which illness data are recorded. The canvass in each city required from 3 to 4 weeks. The dates of the canvass were slightly different in each locality, but fell between March 20 and May 15, 1933, for all localities

Illnesses were classified according to whether their time of onset was within the survey period of 3 months or prior to the survey, the

⁶ All illness rates are adjusted for age, using the method of expected cases as outlined by Raymond Pearl in Medical Biometry and Statistics, pp. 265-269, second edition, 1930. The standard age-specific rates which are used in the adjustment process are rates for all economic groups in all surveyed localities.

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latter including illnesses that were more or less chronic. Each of these 2 groups was further subdivided into disabling and nondisabling cases. All bed cases are included in the disabling class. A disabling illness, whether its onset was within or prior to the survey period, refers to a case causing inability to pursue the usual work, school, or other activities for 1 or more days during the 3 months of the study; 86 percent of the disabling cases with onset within and 69 percent of those with onset prior to the survey were also in bed for 1 or more days during the study period.

ILLNESS EARLY IN 1933 AND UNEMPLOYMENT IN 1932

In table 5 the incidence of illness is shown for 3 groups of the entire surveyed population in the 10 localities classified according to employment status of the wage earners in 1932. Illnesses are shown as (1) All

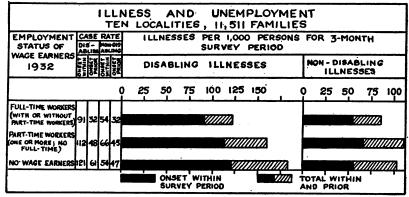


Figure 1.—Incidence of disabling and nondisabling illness in 10 localities during a 3-month period in the early spring of 1933 in white wage-earning families classified according to number of employed workers in 1932. (Rates are adjusted for age.)

cases; (2) nondisabling cases; and (3) disabling cases (a) not in bed, In figure 1 disabling and nondisabling cases are shown for the same groups of the surveyed population as appear in table 5. The chart shows a lower incidence of disabling illness among families having full-time workers than in families having part-time workers only or families having no wage earners. The group with no employed workers has an incidence of disabling illness, onset within the survey period (121 cases per 1,000 persons), that is 33 percent higher than the rate of the group having full-time workers (91 per 1,000). Illnesses with onset prior to the period (largely chronic) are nearly twice as high in the group without employed wage earners as in the group having full-time workers (61 as against 32 disabling cases per 1,000 persons). Combining disabling illnesses having onset within and prior to the study, the unemployed group shows a rate (182 cases per 1,000) 48 percent higher than the families having full-time workers (123 per 1,000). Nondisabling cases with onset within the survey period show no logical relationship to employment status;

TABLE 5.—Illness and unemployment

[Incidence of disabling and nondisabling illness in the early spring of 1933 in 11,511 white wage-earning families classified according to employment status of wage earners during 1932, in 10 localities]

	Case	rate 1	er 1,000	person	s for 8-1	nonth s	urvey p	eriod		
	01	nset wit	hin per	iod	Or	set pric	or to per	iod	Popu- lation	
Employed workers in the family		Non-	Disa	bling		Non-	Disa	bling	ob- served	
,	Total	disa- bling	Not in bed	In bed	Total	disa- bling	Not in bed	In bed		
Full-time workers (1 or more, with or without part-time)	145 178	54 66	13 15	78 97	64 93	32 45	9 15	23 33	21, 02 2 21, 22 4	
No employed workers	175 54				108	47	21	40	4, 935	
Total population 3	163	59	14	90	81	39	13	29	47, 181	

DISABLING ILLNESS AND UNEMPLOYMENT

	יכוט			LLNE								
EMPLOYMENT STATUS OF WAGE	-	S INDEX	4 20	_do ^{IN}	IDEX 60	OF 1		LING	ILLN	ES\$ 160	180	200
EARNERS 1932		CHESET			<u>~~</u>	<u>مح</u>	100	-50	140	<u> 'Y'</u>	<u> </u>	~~
		33 33	BALTI			7773						
FULL-TIME	50 65						****					
NO WAGE EARNERS												
HO MAGE ENKNERS	68	60		10114								
FULL-TIME	71	24	BIRMI	NGHAI	٩	/////	73					
PART-TIME	66	34					=		4			
NO WAGE EARNERS	71	41					7.					
NO MADE CARRERS	''		BROOK	VN				4				
FULL-TIME	74	11	DROOK	LYN								
PART-TIME	114	16				لنند			,			
NO WAGE EARNERS	156	30										
ENKAERS	.30	30	CLEVE	ANC								
FULL-TIME	68	19	VLEVE.	-ARU	Z	<i>i</i> zza	,					
PART-TIME	72	27					772					i
NO WAGE EARNERS	82	42										
HO HADE ENKHERS	06		DETROI	Υ								
FULL-TIME	50	28	DETROI			7						
PART-TIME	73	28					<i></i>					
NO WAGE EARNERS	83	48					_	*****				
NO WHOL LANGERS	00		NEW YO	NR K		111						
FULL-TIME	67	19	1400	, Kiy	111							
PART-TIME	80	28										ı
NO WAGE EARNERS	116	34					7772	,,,,,,				1
			PITTSB	URGH	-,			11111			•	
FULL-TIME	67	22			1/							ł
PART-TIME	72	27					77					
NO WAGE EARNERS	92	38						7777				
			SYRACI	JSE			www.	مس				ı
FULL-TIME	54	19			/////	1						1
PART-TIME	83	32				111		3				
NO WAGE EARNERS	79	46										1
			GREEN	VILLE				لقف				I
FULL-TIME	54	28				\mathbb{Z}						į
PART-TIME	66	40			1							WITHIN
NO WAGE EARNERS	20	40	////		7							Y PERIOD
			MORGA		T							MIHTIN
FULL-TIME	72	17								will f	MD P	RIOR
PART-TIME	80	25				1111						1
NO WAGE EARNERS	66	18			111	Z						ł

FIGURE 2.—Disabling illness in each of 10 localities, during a 3-month period in the early spring of 1933 in white wage-earning families classified according to number of employed workers in 1932. (Illness rates, adjusted for age, are expressed as an index (100 equals the disabling illness rate, adjusted for age, onset within and prior to the survey period, for the entire canvassed population in the specified city).)

Adjusted for differences in age distribution.
 Excludes 1,955 individuals living on income or pension.

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nondisabling cases with onset prior to the period are 47 percent higher in the group having no wage earners than in the group having full-time workers (47 as against 32 cases per 1,000 persons).

In figure 2 and table 6 similar data are given for disabling illnesses for each of the 10 localities. A disabling illness index (100 equals the disabling illness rate, adjusted for age, onset within and prior to the period, for the entire surveyed population in the specified city) is used in figure 2 instead of the actual rate. This eliminates differences in rates from city to city and shows only the relative variation of the illness rate with employment status of the family wage earners. Actual rates adjusted for differences in age distribution, as well as cases of illness and population observed are given in table 6.

Table 6.—Disabling illness in the early spring of 1933 and employment status of wage earners in 1932 in white wage-earning families in each of 10 localities

									-			
	il 1,00 for	oisab Iness 00 pe 3-m surv perio	per erson onth	s C	ases of illr	disabli less	ing		Popu	lation (bserve	d
Locality	Full time	Part time	Unemployed	Full time	Part time	Unemployed	Income or pen-	Total	Full time	Part time	Unemployed	Income or pen- sion
BaltimoreOnset within	68	88	119	168	180	68		5, 167	2, 572	1, 960	531	104
Onset prior Birmingham	45	55	81	119	106	42		4, 137	2, 342	1, 366	322	107
Onset within	105 35	97 51	104 61	243 83	135 69	34 19						
Brooklyn Onset within	81	125	171	.	95		26	3, 547	2, 295	777	110	365
Onset priorCleveland	12	17	33	178 32	16	19 4	18	5, 080	1,814	2, 015	811	440
Onset within Onset prior	89 25	95 35	109 56	150 49	189 70	93 42	26				-	
Detroit	63	93	105	114	256	101	17	5, 633	1,842	2, 676	933	182
Onset prior New York	36	36	61	65	88	53	15	5, 079	2, 947	1, 423	441	268
Onset withinOnset prior	108 31	130 46	186 55	302 96	182 70	92 21	42 13	F 001				
Pittsburgh Onset within Onset prior	102 33	109 41	140 58	206 82	203 81	113 46	15 19	5, 031	2, 151	1,904	800	176
Syracuse Onset within	74	l	108	142	219	102	12	5, 044	2, 022	1, 914	889	219
Onset prior	26	44	63	55	84	53	29	5, 653	1, 594	3, 986	48	25
Onset prior	110 57	134 82	40 80	180 76	563 277	2 3	1 6					
		123	102	166	409	6	3	4, 765	1, 443	3, 203	50	69
Onset prior	27	38	28	32	99	1	8					
Total, 10 localities 2 Onset within	91			1,849	2, 431	630	181	49, 136	21, 022	21, 224	4, 935	1, 955
Onset prior Total eight large cities 3	33	44	58	689	960		167	38, 718	17, 985	14, 035	4, 837	1,861
Onset within	86 l	106	130	1, 503	1, 459		177 153					

¹ Adjusted for age. Rates are not given for the group living on income or pension, because of the small number of persons included in this group in many of the cities. The average disabling illness rates in the group living on income or pension in the 10 localities are as follows: Onset within period, 89 cases per 1,000; onset prior, 87 cases per 1,000. For the 8 large cities, the corresponding illness rates are, respectively, 102 and 63 cases per 1,000 persons.

¹ Illness rates are simple averages of rates in the 10 localities.

Excludes Greenville and Morgantown. Illness rates are simple averages of rates in the 8 large cities.

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With the exception of Greenville and Morgantown it will be seen that the disabling illness rate of families having no employed workers is consistently higher in each city than that of families having part-time or full-time workers. Inasmuch as most of the families having no employed workers in 1932 had one or more employed workers in 1929, these data are striking evidence of the association between a relatively high rate of disabling illness and loss of employment during the depression, with accompanying loss of income and reduced standard of living.

ILLNESS EARLY IN 1983 AND INCOME IN 1932

When families are grouped according to income in 1932, the same inverse association of illness rates with economic well-being is evident

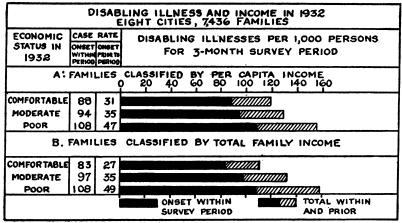


FIGURE 3.—Disabling illness in 8 large cities during a 3-month period in the early spring of 1933 in white wage-earning families classified according to (a) annual per capita income in 1932, and (b) annual total family income in 1932. (Ranges of income included as "comfortable", "moderate", and "poor" are given in footnote 8, page 608. Rates are adjusted for age.)

as in the grouping by employment status of the wage earners. Figure 3 shows the incidence of disabling illness among families in the 8 large cities grouped first according to per capita income and second according to total family income. By either classification the families in the lowest income groups show the highest rates of disabling illness. Thus the rate among families classified as "poor" is 23 percent higher in the grouping by per capita income and 30 percent higher in the

⁷ The 2 rural industrial communities, while having a relatively high average illness rate, do not show the consistent association between economic status and illness which appears in the 8 large cities. This finding, for which there is no obvious explanation at the present time, has made it seem best to consider the large cities as a group for many tabulations and reserve the 2 rural communities for separate study.

grouping by total family income than the illness rate of families classified as "comfortable." Illnesses with onset prior to the period, largely chronic, show an even greater excess among families with the lowest income. Thus the poor group has an illness rate 50 percent higher than the comfortable group in the classification of families by per capita income and 80 percent higher than the comfortable group in the classification by total family income.

Income classification

· · · ·	Ann	ual per capita incom	• -
Cky	Comfortable	Moderate	Poor
I. Baltimore, Birmingham, Cleveland, Detroit, Pittsburgh, and Syracuse. II. Brooklyn and New York City. III. Greenville and Morgantown.	\$425 and over \$500 and over \$300 and over	\$150-\$424 \$250-\$499 \$150-\$299	Under \$150. Under \$250. Under \$150.
au.	Annu	al total family incom	10
City	Comfortable	Moderate	Poor
I. Baltimore, Birmingham, Cleveland, Detroit, Pittsburgh, and Syracuse.	\$1,600 and over	\$600-\$1,599 \$1,200-\$1,999	Under \$600. Under \$1,200.

This excess was not evident in the crude rates which were used in preliminary publications. The adjusted rate for illnesses having onset prior to the study period among the comfortable group is considerably lower than the crude rate, due to the fact that this group includes a relatively large proportion of older individuals with a high rate of chronic illness. Hence, with the effect of differences in age composition elimnated, the "poor" are shown to have a much higher rate of chronic illness than the "comfortable."

^{*} For convenience, incomes have been grouped into ranges classified as "comfortable", "moderate", and "poor." These terms have no significance other than as convenient labels for use in discussion. The income ranges included in these groups are not the same for each city, due to differences in the averages and distributions of the incomes and the necessity for having groups of sufficient size for statistical significance. New York and Brooklyn, for example, had relatively few families with incomes under \$000, and the "poor" group in those cities includes all families with incomes under \$1,200. The need for the change in income class limits for certain of the localities is also indicated by higher and lower living costs in the communities concerned. Per capita income has been used in many of the tabulations because it represents economic status better than the total family income which takes no account of size of family. It was realized that for strict accuracy a figure taking account not only of the size of the family but also of the age and sar of its members, such as "income per adult male unit", might be better than income per capita. However, previous studies have shown excellent correlation between per capita income and these other derived units, and it was felt that the accuracy of the 4-year income record was not sufficient to justify the more refined calculations. The income ranges used in all charts and tables are as follows:

TABLE 7.—Disabling illness in the early spring of 1938 and family income 1 in 1932 in white wage-earning families in each of 10 localities

Disabling	Ξ	lness p	illness per 1,000 persons for 3-month survey period :	ersons fo	or 3-mon	th sur-		Cases	Cases of disabling illness	alli gail	888		-		Popule	Population observed	erved		
	15. S	lassified by pe capita income	Classified by per capita income	Clas	Classified by total family income	total	Classicapi	Classified by per capita income	per 10	Classifi famil	Classified by total family income	-	Total	Class	Classified by per capita income	per ne	Class	Classified by total family income	otal 16
	e te	om- rt- ole erate	d- te Poor	Com- fort- able	Mod- erate	Poor	Com- fort- able	Mod-	Poor	Com- fort- able	Mod- erate	Poor	popu- lation	Com- fort- able	Mod- erate	Poor	Com- fort- able	Mod- erate	Poor
	۱ : ۱	<u> </u>	L		<u> </u>							Ļ	4,442	381	2, 261	1,800	\$	2,544	1,24
54.0	-4	878	72 94	39.7	22	5 6	58	129	38	4 %	282	- 168 - 168	-	1					
1	٠,	- 1	-	_;	-	}				1	1	<u>: </u>	3,348	3	1,561	1, 233	ş	1,612	1,081
8.4 8.4	20.4	284	30	88	==	32	42	2 <u>8</u> 2	1 <u>2</u> 2	28	178 2	සි :	Ť	Ť					
•	`				- :		1	3	3		3	<u>:</u>	2,703	879	1, 185	88	842	86	88
28.	~ -	85	13 110	282	35	118	\$5	108	25	8:	8 =	85	1	i				Ì	
	` [-	_;		_	3	1	1	3	2		1	4, 415	425	1,690	2,300	513	1,891	2,011
98.8	- س	25.5	91	388	32	29	38	<u> </u>	2 8	85	22	- 2 2 3 3	†	Ī	Ī			Ī	
0	1			-	_	a.	1	7	8	3	8	<u>:</u>	4, 555	452	1,809	2, 204	8	2,066	1,880
200 200 200 200 200 200 200 200 200 200	.→6	25	88	28 s	28	83	83	159	200	3 :	162	187	1						
	•					2	9	1	3	1	2	2	. 541	626	1.590	2.022	701	1.536	2.305
122.4	≍'	200	160	102	116	132	25:	165	283	\$	172	8	Ť						
3	•	_					₽	3	3	 7	20	÷	2 480	KOR	1 274	1 500	A2K	1 689	1
110.6	-	- 28	99 12	127 98	98	135	\$	129	201	99	7	173	3	3	1,014	7, 00	3	3	7, 410
 						8	27	22	2	8	6	÷	171	200	1 820	9 212	462	210	1 700
101.5	=	8	96	2 78	106	104	25	140	246	32	508	170		3	1, 006	6,010	3	4	3
		37					8	8	88	=	12	<u>.</u>	4.714	354	1.456	2.904	621	1.933	2 160
122.4	=	122	111 12	78	134	124	37	163	405	9	267	292							
7.0							22	103	<u> </u>	43	121	•	3 835	302	918	2.618	402	1 479	1 871
114.5	: ="	458	127 10	106 121	101	228	8 ,	121	301	8°	167	82		3	3				
		ļ	1	1	1	ļ	1	3	1	°	+	÷	101	100	18 970	13	300 0	14 610	070 01
0	-	g	<u>:</u>	!	<u>;</u>	1	441	÷	9 261	609	1 747	_	10, 10	97	10, 3/3	30/ AT	0, 270	ATO '/I	10,04
4	. •••	8 %	88	40	8	12	8	., 64.	872	క్షే	202	918	31 635	4 482	13 002	14 181	K 112	14 214	12 200
0.8		8 =	26	108	26	802	358	1, 165	1,555	394	1, 323	1,381		,					
	1							ı			١								

1 For definition of the groups "comfortable", "moderate", and "poor", see footnote 8, page 608.
3 Adjusted for age.
3 Illness rates are simple averages of rates in 10 localities.
4 Excludes Greenville and Morgantown. Illness rates are simple averages of rates in 8 large cities.

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In table 7 disabling illness rates are given for each of the 10 localities for families classified by per capita and by total income. In figure 4 for families classified by per capita income a disabling illness index (100 equals the disabling illness rate, adjusted for age, onset within and prior to the period, for the entire surveyed population in

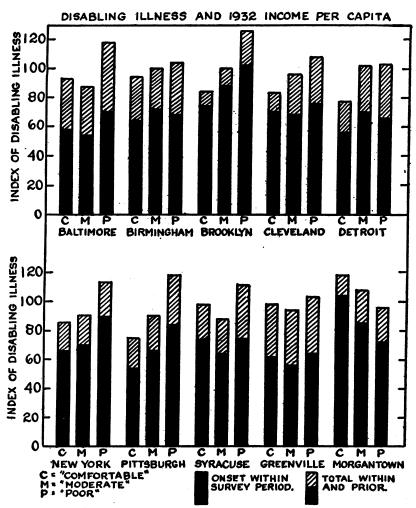


FIGURE 4.—Disabling illness in each of 10 localities during a 3-month period in the early spring of 1933 in white wage-earning families classified according to annual per capita income in 1932. (Illness rates, adjusted for age, are expressed as an index (100 equals the disabling illness rate, adjusted for age, onset within and prior to the survey period, for the entire canvassed population in the specified city. Ranges of income included as "comfortable", "moderate", and "poor" are given in footnote 8, page 608).)

the specified city) has been used instead of the actual rate. Considering illnesses having onset within and prior to the study period, sickness rates in the poor group (by per capita income) are consistently higher than in the comfortable group, with the exception of Morgantown. In the classification by total income, Morgantown

shows the same association with economic status as the other localities, the lowest income class having the highest sickness rates.¹⁰

ILLNESS EARLY IN 1933 AND INCOME CHANGE, 1929-1932

A correlation between sickness and low income is not confined to periods of depression. A high illness rate, high death rate, and high birth rate have always gone hand in hand with poverty. It is obviously desirable, therefore, to ascertain whether the higher sickness rate among the poorer classes in the surveyed families was in any way associated with changes in standard of living. Tremendous shifts in economic status and standard of living took place during the depression. For example, of the 14,181 individuals in the eight large cities who were classified by per capita income as poor in 1932, only 25 percent were poor in 1929, 55 percent were moderate, and 20 percent were comfortable. An analysis of the relation between "depression history" and illness was made. For this purpose the individuals were divided into six categories according to economic status in 1929 and 1932, as follows: 12

- I. Individuals experiencing materially lowered family income between 1929 and 1932 were classified as—
 - 1. Comfortable in 1929 and moderate in 1932.
 - Moderate in 1929 and poor in 1932.
 Comfortable in 1929 and poor in 1932.
- II. Individuals who had not experienced materially lowered income between 1929 and 1932 were classified as—
 - 1. Comfortable in 1929 and 1932.
 - 2. Moderate in 1929 and 1932.
 - 3. Poor in 1929 and 1932.

Sickness data for these groups classified according to per capita income are given in figure 5. Inspection of the chart shows the significant and interesting fact that the highest illness rate is exhibited by the group hardest hit by the depression, namely, the group "comfortable in 1929 and poor in 1932." Considering disabling illnesses having onset within or prior to the survey period, this group,

¹⁸ If the differences in illness rates between the comfortable and poor groups in the individual localities are tested for statistical significance, it is found that the differences are from 1 to 4 times their respective probable errors, which vary from 10 to 14 cases per 1,000 persons in the several localities. Thus in Birmingham and Syracuse, where the difference in illness rates (onset within and prior) between the comfortable and poor groups is 17 and 15 cases per 1,000, respectively, the association between economic status and illness is within the limits of chance variation. However, the probability of finding a consistent association between income and sickness in this number of cities, as a result of chance, is so small that the relation is unquestionably real. This applies also to the differences in illness rates observed among families grouped by employment status of wage earners (table 6) or by change in income between 1929 and 1932 (tables 8 and 9). Considering the average results for the 8 large cities, the poor group exhibited a rate of disabling illness, onset within and prior, which was 36 cases per 1,000 above that of the comfortable group. The probable error of this difference is 4 cases per 1,000; thus the actual difference observed is 9 times its probable errors.

[&]quot;See Public Health Bulletin 165, Economic Status and Health (Govt. Printing Office, Wash., 1927), for a summary of data bearing on the association of illness and death rates with economic status.

¹² Ranges of income included as "comfortable", "moderate", and "poor" are given in footnote 8, p. 608.

with a rate of 174 cases per 1,000 persons, showed an incidence of illness that was 45 percent higher than the rate (120 per 1,000) for their more fortunate neighbors who were equal in status in 1929 but suffered no drop in income by 1932; that is, the "comfortable in 1929 and 1932." ¹³ The group that had dropped from comfortable to moderate showed a 10 percent higher disabling illness rate than the comfortable group that had experienced no drop in income. The group that had dropped from moderate to poor showed a 17 percent higher illness rate than those who were in moderate circumstances throughout the 4 years. It is interesting to note that the rate for

DISA	BLING II	LLNE	55	AND CHANGE IN PER CAPITA INCOME
			1	EIGHT CITIES
ECONOMIC	STATUS	CASE ONSET		DISABLING ILLNESSES PER 1,000 PERSONS FOR 3-MONTH SURVEY PERIOD
1929	1932	SURVEY PERIOD	TO	20 40 60 80 100 120 130 140
	PERSON	s w	ITH (DIMINISHING INCOME, 1929-1932
COMFORTABLE	MODERATE	97	35	
MODERATE	POOR.	103	42	
CONFORTABLE	POOR	121	53	
1	I PERSON	s w	ITH L	UNCHANGED INCOME, 1929-1932
COMFORTABLE	COMFORTABLE	90	30	WIIIIII.
MODERATE	MODERATE	90	34	
POOR.	POOR	107	52	
				ONSET WITHIN TOTAL WITHIN AND PRIOR

FIGURE 5.—Disabling illness in 8 large cities during a 3-month period in the early spring of 1933 in white wage-earning families classified according to change in per capita income, 1929-1932. (Ranges of income included as "comfortable", "moderate", and "poor" are given in footnote 8, page 608. Rates are adjusted for age.)

the group that had dropped in income from comfortable to poor was 9 percent higher than that of the chronic poor, that is those who were poverty stricken even in 1929—a finding which suggests that illness is associated with sudden change in standard of living.

¹⁵ In preliminary tabulations a larger number of income groups was used, each group including a narrow range of incomes. It was found, however, that the broad groups finally used were adequate. For example, the "comfortable" class (\$125 and over by per capita income) was divided into 3 groups, (1) \$425-\$499, (2) \$500-\$749, and (3) \$750 and over. It was found that the illness rates among families that had dropped in income from either of these classes into the "poor" group were similar and were all higher than in families that remained in either of the three classes from 1929 to 1932. Similar subdivision of the "moderate" and "poor" groups was made and found not to change the general picture as presented in this paper.

Table S.—Disabling illness in the early spring of 1933 and change in annual per capita income, 1929-32, in white wage-earning families in each of 10 localities

							ence of to cocures	77 60 1) tocate	9277									
	Disal	oling ill 3-m	Disabling illness per 1,000 persons 3-month survey period?	r 1,000 rvey pei	person:	s for		Cases	Cases of disabling illness	oling illa	esa				Popula	Population observed	rved		. ,
Per capita income: 1 1929.	Com- fort- able	Mod- erate	Comfort.	Com- fort- able	Mod- erate	Poor	Com fort.	Mod- erate	Com- fort.		Mod- erate	Poor	3	Confort.	Mod- erate	Com- fort- able	Com- fort- able	Mod- erate	Poor
1932	Mod- erate	Poor	Poor	Com- fort-	Mod- erate	Poor	Mod- erate	Poor	Poor	Com- fort- able	Mod- erate	Poor	 1500 T	Mod- erate	Poor	Poor	Confort.	Mod- erate	Poor
Baltimore												İ	4.374	895	88	340	381	1.328	374
Onset within	6. 5.	80	711	84	67	88	84	뒃꼆	#8	88	82	38		1	+				
Birmingham	2	3	2	2	2	5	1	3	1	1	3	<u>-</u>	3, 316	8	82	322	536	288	191
Onset within	123	88	118	84	828	22	116	3.2	38	2 %	£ 4	4.5							
Brooklyn	2	2	3	2	3	:		•		1	:		2,565	572	327	162	783	571	251
Onset within	127	01:	88	29	8;	162	E,	8	21	22	8	87	-	-	+	i	1	+	
Cloveland	3 _	2	ŝ	3	or	8	•	0		=	2	•	4.333	814	1.179	467	402	817	9
Onset within	88	85	128	87	g	Ξ	69	102	8	31	2	72							
Onset prior	41	34	49	18	32	22	\$	33	8	2	8	88	437.7	790	036	8	- 417	- 100	777
Detroit mithin	4.0	64	9	13	9	69	1	119	84	g	5	g	Ž.	3	A .	3	21	8	\$
Onset prior	2 %	36	2 4	28	§ 4	4	5.4	18	3 25	3 =	28	38					-		
New York City.	3	3	1	1	:	3							4, 411	731	888	334	846	812	749
Onset within	117	142	138	106	20:	140	82	138	∞	38	28	96	-		Ì		1	+	:
Dittshurgh	8	4.	3	70	7	ò	5	ò	0	ò	8	5	3 373	681	3	307	883	620	302
Onset within	84	119	176	8	108	108	75	107	S	9	99	7							
Onset prior	88	22	22	35	35	45	8	4	ຊ	8	ន	15	200	000	040	8	-	-	
Syracuse	- 1	1	!	-	8			199		- 00	-		, 9	200	1, 0/0	8	- 5	150	ž
Onset prior	38	47	130	38	32	125	88	32	28	5 52	28	₹8							
Greenville	-	!									-	į	4, 624	440	1, 392	22	317	229	28
Onset within	38	131	F 5	132	26	35	3 2	2,2	3	88	2		-					-	:
Morgantown	•		3	3	6	3	3	2	3	}		3	3, 797	649	1,042	250	28	8	828
Onset within	109	Ξ8	112	<u> </u>	151	48	12	328	88	38	¥,€								
Total, 10 localities 3	:	3	1									-	39, 337	7.833	10, 208	4.328	4. 725	7.076	5.174
Onset within	102	107	121	88	96	107	783	1, 133	828	410	614	570						Ī	
Total. 8 large cities	8	Ç.		Ş	SS .	3	8	ş	077	181	3	÷.	30, 916	6,340	7, 769	2,820	4, 117	6, 278	3, 592
Onset within Onset prior	35	103	123	88	8%	107	283	2862	356	88.40	222	392	$\overline{\parallel}$						
For definition of the ground	1 0	mfortab	., ", "le	noderate	a", and	"poor"	"comfortable", "moderate", and "noor", see footnote 8, D.	tnote 8.	. p. 608.										

For definition of the groups "comfortable", "moderate", and "poor", see footnote 8, p. 608.

Adjusted for age.

Excludes 47 persons in families with rising income, 1929-32. Illness rates are simple averages of rates in the 10 localities.

Excludes 847 persons in families with rising income, 1929-32. Illness rates are simple averages of rates in the 8 large cities.

Excludes Greenville and Morgantown. Excludes 719 persons in families with rising income, 1929-32. Illness rates are simple averages of rates in the 8 large cities.

Table 9.—Disabling illness in the early spring of 1933 and change in annual total family income, 1929-32, in white wage-earning families

						ı	s each	01 TO 1	in each of 10 localities	•								
	Disal	Disabling illness per 1,000 persons 3-month survey period ?	ness per	1,000 vey per	persons od ?	for		Cases of	Cases of disabling illness	g illnes	8			Popula	Population observed	rved		
Total family income: 1929.	Com- fort- able	Mod- erate	Com- fort- able	Com- fort- able	Mod- erate	Poor	Com- fort- able	Mod- fo	Com- Com- fort- fort- able able	The Mod-	Δ,	of Total	Com- fort: able	Mod- erate	Com- fort.	Com- fort- sble	Mod- erate	Poor
1932.	Mod- erate	Poor	Poor	Com- fort-	Mod- erate	Poor	Mod- P	Poor P	Poor fort-	Com- fort- able erate	d- te Poor			Poor	Poor	Confort.	Mod- erate	Poor
Baltimore												4.356	1.187	787	2	019	1.315	118
Onset within	\$:	25	113	83	25	61	83	8:	88	38	125	=======================================	÷			-	+	
Birmingham	6	7.7	70	\$	ì	۵,	3	8	7		200	3, 299	1.131	202	364	883	454	160
Onset within	125	117	8:	æ:	23:	8:	141	88	ຂ:	22	37	01	÷			-	+	
Unset prior Brooklyn	2	8	‡	3		્	2	25	9		-	2.580	470	420	237	792	187	196
Onset within	88	101	88	78	8	146	\$	46	ຂ	82	45	27					i	
Onset prior	. 15	8	=	2	=	8	x 0	2		<u>.</u>	9	8 4 205	1 010	1 0/4	603	757	808	354
Onset within	98	92	101	52	103	123	16	8	69	55	-62		_	3	3		3	5
Onset prior	8	3	25	8	8	29	8	4	33	12	:8	Z	! ;					
Detroit			-			-		:	-	-	-	4, 461	1,277	877	748	220	3 5	727
Onset within	28	8 \$	2.5	3,2	88	Z 2	28	825	≅ 8	25	88	28	-				+	
New York	3	2	3	3	5	:	-	3	3	1	3	4.387	629	934	208	229	83	8
Onset within	38	85	117	88	130	145	555	125	89	8	8	116	4				i	
Pittshingh	38	35	\$	38	;	‡	29	3	17	R	3	8 871	878	760	254	102	620	150
Onset within	88	128	174	8	- 26	8	12	8	8	49	20			3	3	3	3	
Onset prior	42	28	22	æ	32	\$	7	7	2	22	S	90	-				-	
Syracuse				-	-	-		- :		:	-	4,034	893	972	8	416	1,022	8
Onset within Onset prior	8 7	55.88	ន្ទន	ន្តន	118	<u>5</u> 2	328	53	3 2	===	35 25 25 26 27 27 27 28 27 28 28 28 28 28 28 28 28 28 28 28 28 28	25 SE	- <u>!-</u> ;			1		
Onset within	131	130	119	8	136	108	167	170	56	45	95	30.	1,2,1	1, 100	Ì	3	5 0	3
Onset prior	8	2	22	7.	22	118	1	2	2	8	3	8	<u> </u>					
Morgantown		200						-	8	-		3, 712	2 1,059	88	694	83	88	8
Onset prior	38	34	46	217	<u> </u>	32	31	38	38	5 8 0	80	201						
Total, 10 localities 3											<u> </u>	39, 140	0,804	8,310	5,086	5,745	7, 251	2.94
Onset within	88	11.5	112	88	<u>\$</u>	85	88	974	281	3 2	986	347	÷	÷				
Total, 8 large cities	3	3	-	3	3	5	1	3	-	-	-	30, 783	3 7.484	6,320	3,642	4, 716	6, 274	2,247
Onset within Onset prior	28	88	55	88	82	202	28 88 88	88	397 159	357	630 213	275						
1 For definition of the grou	2	"comfortable", "moderate",	m, ,, m	nderate	1.	and "noor"	see footnote 8.	note R. n	8									

For definition of the groups "comfortable", "moderate", and "poor", see footnote 8, p. 608.
 Adducted for age.
 Excludes of rage.
 Excludes 1.04 persons in families with rising income, 1929-32. Illness rates are simple averages of rates in the 10 localities.
 Excludes Greenville and Morgantown. Excludes 852 persons in families with rising income, 1929-32. Illness rates are simple averages of rates in the 8 localities.

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In figure 6, the results for each of the 10 localities are shown for 2 economic groups classified by per capita income in 1929 and 1932, (a) comfortable in 1929 and 1932, (b) comfortable in 1929 and poor in 1932. With the exception of Greenville, a higher illness rate is exhibited in each locality by the group that had dropped from comfortable to poor than by the one that remained in the comfortable

	DISA	BLIN	G IL	LNESS AND CHANGE IN PER CAPITA INCOME
ECONOMIC	STATUS	ILLNES		20 40 60 80 100 120 140 160 200
1929	1932	ONSET WITHIN	ONSET PRIOR	BALTIMORE
FORTABLE	COM- FORTABLE	60	33	
FORTABLE	POOR	86	58	
				BIRMINGHAM
COM- FORTABLE	COM- FORTABLE	. 64	30	
FORTABLE	POOR	81	35	
				BROOKLYN
COM- FORTABLE	COM- FORTABLE	73	9	<i>W</i> .
COM- PORVABLE	POOR	56	36	
				CLEYELAND
FORTABLE	COM- FORTABLE	65	14	
FORTABLE	POOR	96	36	
				DETROIT
FORTABLE	COM- FORTABLE	59	23	
FORTABLE	POOR	79	38	
				NEW YORK
COM- FORTABLE	COM- FORTABLE	68	20	
COM- FORTABLE	POOR	87	15	
				PITTSBURGH
COM- FORTABLE	COM- FORTABLE	57	22	
COM- FORTABLE	POOR	114	37	
				SYRACUSE
COM- FORTABLE	COM- FORTABLE	77	22	
FORTABLE	POOR	.94	53	
				GREENVILLE
COM- FORTABLE	COM - FORTABLE		32	ONSET WITHIN SURVEY PERIOD
FORTABLE	POOR	64	33	TOTAL WITHIN
		L		MORGANTOWN AND PRIOR
COM- FORTABLE	COM- FORTABLE	84	15	
COM- FORTABLE	0000	75	32	

FIGURE 6.—Disabling illness in each of 10 localities during a 3-month period in the early spring of 1933 in white wage-earning families classified as "comfortable in 1929 and 1932" and "comfortable in 1929 and poor in 1932". (Illness rates, adjusted for age, are expressed as an index (100 equals the disabling illness rate adjusted for age, onset within and prior to the survey period, for the entire canvassed population in the specified city). Ranges of income included as "comfortable" and "poor" are given in footnote 8, page 608.)

class for the 4 years. In table 8 disabling illness rates are given for all of the economic groups classified by 1929 and 1932 income per capita; and in table 9 illness rates are given for families grouped by total income in 1929 and 1932. Classification by total family income gives, in general, the same sequences as classification by per capita income.

ILLNESS EARLY IN 1983 AND RELIEF STATUS IN 1982

In 1932, in the 8 large cities 20 percent of the surveyed families received public or private relief for all or part of the year. The proportion on relief varied from 4 percent in Brooklyn to 30 percent in Syracuse (table 1). At that time (1932 and 1933) eligibility for relief indicated that a family was in very dire straits. These relief families had the lowest standards of living of any in the surveyed group. It will be of interest to compare their illness record with that of families not on relief.

Relief families were nearly all in the group classified as poor in 1932 (footnote 8, p. 608). Hence only this group has been separated into relief and nonrelief classes. In figure 7, rates of disabling illness are shown for individuals classified by economic status in 1929 and 1932

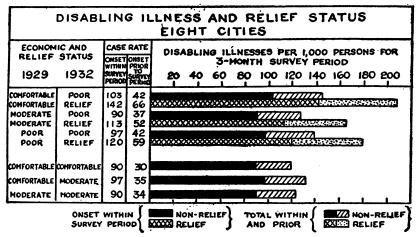


FIGURE 7.—Disabling illness in 8 large cities during a 3-month period in the early spring of 1933 in white wage-earning families classified according to change in per capita income, 1929-32, and relief status in 1932. (Ranges of income included as "comfortable", "moderate", and "poor" are given in footnote 8, page 608. Rates are adjusted for age.)

with the groups that were poor in 1932 classed as (1) poor but not on relief and (2) poor and on relief. It is seen that individuals in families on relief have a higher incidence of disabling illness than any of the other groups of the surveyed population, whatever their economic history during the depression. Thus, the group that dropped from the comfortable class in 1929 to relief in 1932 exhibits an illness rate (within plus prior) 44 percent higher than that of the group that fell from comfortable to poor but not on relief and 73 percent higher than that of the group that was comfortable in 1929 and 1932. Among relief families, the income change between 1929 and 1932 is associated with illness in the same manner as for families not on relief; that is, the families that suffered the greatest change in economic status exhibit the highest illness rate.

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In figure 8, illness rates for each of the 8 large cities are shown for 3 groups of families: (1) Comfortable in 1929 and 1932; (2) comfortable in 1929 and poor in 1932; and (3) comfortable in 1929 and on relief in 1932. To facilitate comparisons, a disabling illness index is used instead of the actual illness rate. With the exception of Brooklyn and Birmingham, the highest illness rate is shown by the group that was comfortable in 1929 but on relief in 1932. In Brooklyn the group on relief was too small to give illness rates of statistical signifi-

	DISABL	ING	ILI	LNESS AND RELIEF STATUS.
ECONOM RELIEF	STATUS	WITHIN	ONSET PRIOR	I INDEX OF DISABLING ILLNESS
1929	1932	SURVEY PERIOD	SÚŘVEY PERIOD	BALTIMORE
CONFORTABLE	COMFORTABLE	60	33	VIII III III III III III III III III II
	POOR	31	35	
н	RELIEF	114	73	
				BIRMINGHAM
COMFORTABLE			30	
"	POOR	81	29	
- 14	RELIEF	75	56	777
				BROOKLYN
COMFORTABLE			9	<u>//</u>
0"	POOR	59	36	
69	RELIEF	35	_	
				CLEVELAND
COMFORTABLE			13	
"	POOR	105	23	
N	RELIEF	90	46	
				DETROIT
COMFORTABLE			23	
	POOR	71	27	VIIII)
H	RELIEF	86	51	
				NEW YORK
COMFORTABLE	COMFORTABLE		20	7///
u	POOR	82	16	V//)
H	RELIEF	97	12	///
				PITTSBURGH
COMFORTABLE			23	
11	POOR	70	42	
18	RELIEF	152	34	
L				SYRACUSE
COMFORTABLE			23	
	POOR	65	52	
•	RELIEF	1112	54	
	CHSET WIT	HIN S	URVE	PERIOD TOTAL WITHIN AND PRIOR

FIGURE 8.—Disabling illness in each of 8 localities during a 3-month period in the early spring of 1933 in white wage-earning families classified as "comfortable" in 1929 and (1) "comfortable", (2) "poor," and (3) "on relief" in 1932. (Illness rates, adjusted for age, are expressed as an index (100 equals the disabling illness rate, adjusted for age, onset within and prior to the survey period, for the entire canvassed population in the specified city). Ranges of income included as "comfortable" and "poor" are given in footnote 8, page 608.)

cance. In the other cities except Baltimore the group comfortable in 1929 and poor but not on relief in 1932 exhibits a lower illness rate than the relief group but higher than the group which was comfortable in 1929 and 1932. In all of the 8 cities except Baltimore the group which was comfortable in 1929 and poor but not on relief in 1932 has a higher illness rate than the class which was comfortable in 1929 and 1932. Results for the relief and nonrelief groups are given in detail in table 10.

Table 10.—Disabling illness in the early spring of 1933 as associated with change in annual per capita income, 1929-52, and relief status, 1911.

Per capita income and	Disabli	ng illnes	ses per 1 survey	ling lilnesses per 1,000 persons for 3-month survey period ³	ons for 3-	month		Cast	ss of disa	Cases of disabling illness	less			Po	pulation	Population observad	-	
relief status: 1929	Comfo	ortable	Mod	Moderate	Poor	, o	Comfortable	rtable	Moderate	srate	Poor	*	Comfortable	table	Moderate	rate	Poor	*
1932	Poor	Relief	Poor	Relief	Poor	Relief	Poor	Relief	Poor	Relief	Poor	Relief	Poor	Relief	Poor	Relief	Poor	Relief
Baltimore Onset within	27	155	8/	8	69	102	101	8	£3.	19	=	8	ឌ	218	230	292	8	214
Dirmingham	\$	3	\$	8	33	9,	2	73	3	8	*	2	230	2	447	272	6	100
Onset within Onset prior	42	109 81	§ 2	228	27	38	8=	30	88	នដ	00	200	8	*	100	15	98	9
Onset within	40°S	88	112 18	8	154	214 70	3 0	-	32	m	ಷ _ಣ	₩	016	250	\$	2652	ş	233
Onset within Onset prior	140 31	82	4 28	83	108 88	818	81-	32 12	£ 81	20 10 10	28	38	284	Ş	92	Ş	ē	186
Onset within Onset prior	88	107 88	38	<u>.</u> 3	33	4.83	80	82	30 17	28	00 F-	83	\$ 6	3	Ş	8	\$	Ş
Onset within Onset prior Pittsburgh	82.53	153 19	22.4	28 8 8 8	818	3 3	ජික	22 CR	22.22	88 13	3 ::	52	9	2	3 5	2	218	1.1
Onset within Onset prior	82.2	22	115 38	122	83	15 88	11	æ အ	47 16	88	7°	20	2	183	8	845	8	
Onset within Onset prior	9 8	162 79	35 38	113	28	121 71	10	31 16	នដ	28	52 to	\$ 8		1	3	3	}	
Total, 8 cities 3 Onset within	103	142	90	113	97	120 58	153	202	381	424 152	167 86	22.0	1, 479	1, 328	4, 226	3,516	1, 761	1, 806

1 Feet definition of the groups "comfortable", "moderate", and "poor" see footnote 8, p. 608.
2 Adjusted for age.
Wpigined average.

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DISCUSSION OF RESULTS

The general result is clearly shown, by surveys of samples of the poorer sections of eight large cities, that wage-earning families reduced to poverty during the depression suffered to a greater extent from disabling illness in 1933 than their more fortunate neighbors. Individuals in families supported by public or private relief exhibited a higher illness rate than any other group. This finding was true for children as well as for adults and in general for respiratory and non-respiratory illnesses, with the exception of the communicable diseases of childhood.¹⁴ Whatever the implications of the results, the fact remains that illness was most prevalent among those who could least afford this handicap.

However, the survey data raise the question of the relative importance of nurture and nature in bringing about the observed results. In other words, did reduced standard of living cause *increase* of illness among the new poor between 1929 and 1933 or were they more sickly than their neighbors even in 1929? Have we observed the *effect* of the depression on health or merely the results of a great sifting process?

In considering factors that may have brought about the situation in which a group of families characterized by a newly acquired poverty reported a relatively high illness rate, the methodology of the survey must be borne clearly in mind. All sickness data are for a 3-month period early in 1933 with no data for 1929 or other years; the economic data cover the years 1929 to 1932. If we find, as has been shown, a higher illness rate among the depression poor than existed among families remaining in the comfortable class for all 4 years, then it seems reasonable to suppose that reduced standard of living, including crowded housing conditions and lack of adequate food and clothing and medical care, which accompanied this loss of income, had a part in causing this higher sickness rate in 1933.

However, other factors may have played a part:15

(1) Unemployment of wage earners due to sickness probably contributed to the loss in income of certain families; these persons may have been concentrated in the group that suffered economic reverses during the depression and have been responsible for at least a part of the high illness rate in this group. However, analysis of the data shows this to be a relatively unimportant factor. Individuals unemployed due to sickness were not concentrated among the new poor, and, furthermore, the same excess in sickness rates was observed in this group when all families were excluded in which there was unem-

¹⁴ A forthcoming paper will analyze the results by age and by type of illness.

¹⁸ Knowingly false or unconsciously exaggerated reporting of illness by the poorer groups of the population does not appear to be a factor in the results observed, because the observed variation of illness with age, sex, and diagnosis agrees with other known data. Only an omniscient housewife could invent this complicated pattern.

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ployment due to sickness at any time between 1929 and 1932 (prior to the survey period).

(2) The depression may have been a sifting process, separating the fit from the unfit. In spite of innumerable exceptions, the men who kept their jobs were, on the average, the more vigorous, capable, and intelligent ones. Moreover, with many exceptions, those who lost their jobs were less efficient than those who remained employed. This inefficiency may have been exhibited in many ways distinct from inability to compete in the economic struggle-perhaps a diathesis or tendency toward sickliness existed among these families as a concomitant of the economic inefficiency of the wage earner. This explanation of the higher sickness rates among the new poor does not assume sickness per se as a cause of unemployment, but postulates an inherent inferiority of which unemployment was one manifestation and ill health another. According to this hypothesis, the "new poor" would have exhibited a high illness rate even in 1929 (if they could have been singled out for observation), and their lowered standard of living during the depression was not the prime cause of their high illness rate.

The writers admit the possibility that selection played a part in bringing about the situation observed in 1933, but it does not seem probable that selection of the less fit by the depression screen is the whole story. Undoubtedly, those who became unemployed during the depression were, on the average, the least well equipped to compete in the keen struggle for jobs. For example (table 11), when we compare the "new poor" in the surveyed group with those who remained comfortable throughout the depression, we find that they had fewer household heads with high school or college education, fewer in the white-collar occupations in 1929, that they lived in more crowded living quarters even in 1929, and exhibited a higher birth Some of these findings appear to indicate that families of certain types were least successful in weathering the depression. However, it seems highly improbable that a theory of selection contains the sole explanation of the results of the present survey. a matter of fact, when illness rates are made specific for age, sex, race, education, occupation, and relief status, the association between drop in income and high illness rate is still evident.

A study now being made of the death rate among families who became unemployed during the depression will throw further light on the question, because it is possible to obtain information on deaths for a number of years prior to the canvass, which is not feasible in a sickness survey. Hence, trends in the death rate from 1929 to the present time can be studied for groups of families that had various types of economic history during the depression. Preliminarly results indicate a rise in the death rate between 1929 and 1933 among families in which the wage-earner became unemployed during this period.

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TABLE 11.—Characteristics of white wage-earning families classified according to per capita income change, 1929-32: 5 cities surveyed early in 1933 1

	Comfortable in 1929 and 1932 3	Comfort- able in 1929, poor in 1932 3	Poor in 1929 and 1932 2
Percentage of all families:			
With full-time workers, 1929	89. 4	88.3	33. 1
With full-time workers, 1932	72.7	7.0	19. 7
With no employed workers, 1932	. 7	36.8	34. 6
With chief wage-earner in white-collar occupation in 1929	33. 4	9.6	13. 0
On relief, 1929	.0	.6	14. 7
On relief, 1932	.7	55. 9	55. 9
With household head native of native parents	44. 3	43.3	26. 3
With household head having high school or college education	27. 9	19.4	7. 2
With unemployment due to sickness, 1931-32.	6.3	6.0	9. 1
Persons per family, 1933	2.8	4.0	6. 1
Persons per room, 1929	. 54	. 78	1. 2
Persons per room, 1933	. 55	. 93	1. 2
Annual birth rate 2 per 1,000 married women, aged 15-44 years, 1929-			
32	107	133	178
Disabling illness per 1,000 persons for 3-month period 4	119	185	153

Adjusted for age.

The facts that the excess in illness rates appears among children as well as adults and that the highest illness rates are exhibited by families that had dropped from the highest level in 1929 appear to point to a definite causal relation between lowered standard of living and high illness rate. But whatever the cause, the result of the depression has been to present to society for support a group of some 20 million persons in the United States who are on relief rolls and among whom sickness is probably more prevalent than in the rest of the population. It must be recognized that medical care and preventive services for these persons are a necessity of life as well as food, clothing, and shelter. These necessities must be made available to all if the health of the wage-earning population is to be maintained.

SUMMARY

Records of illness during a 3-month period early in 1933 and economic history from 1929 to 1932 have been collected from about 12,000 wage-earning families in the poorer sections of 8 large cities, a group of coal-mining communities, and a group of cotton-mill villages. This paper, the first of a series dealing with the investigation, presents the method of the study and general results for each locality.

Tremendous changes in economic status and standard of living took place among the surveyed families during the depression. The median income of the group in the 8 large cities dropped from \$1,650 in 1929 to \$870 in 1932. In 17 percent of the families the chief wage earner was without employment in 1932; in 10 percent of the families all wage earners were unemployed that year. and private relief agencies contributed to the support of 20 percent of the families for part or all of 1932.

¹ Baltimore, Cleveland, Detroit, Pittsburgh, and Syracuse.
² For definition of groups "comfortable" and "poor", see footnote 8, p. 608.
³ Total family income was used in classifying families for birth-rate tabulation. "Comfort annual family income of \$2,000 and over; "poor", under \$1,200. (Rates adjusted for age.) "Comfortable" indicates

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Disabling illness was found to be 48 percent higher among families having no employed wage earners in 1932 than in families having full-time workers. The group of families that had dropped from fairly comfortable circumstances to relief rolls during the depression showed a rate of disabling illness 73 percent higher than that of their more fortunate neighbors who had remained in the comfortable class throughout the 4 years. The higher sickness rates were observed in general in each of the 8 large cities as well as in the group as a whole. No consistent association between illness and economic status was found in the two rural industrial communities. Insofar as disabling illness is evidence of ill health, the results of the survey show that families hardest hit by the depression suffered to a greater extent from ill health in 1933 than others who had weathered the depression more successfully.

While concentration of the less fit in the ranks of the unemployed may have played a part in bringing about the situation observed in 1933, it does not seem probable that selection is the whole story. Particularly significant are the facts that the highest illness rates were observed among those who had suffered the greatest change in standard of living and that the excess in illness existed among children as well as adults. Whatever the cause, the fact remains that illness was most prevalent among families reduced to poverty and on relief rolls, who could least afford this handicap.

In forthcoming papers analysis of illnesses will be made by cause, by age and sex, and by social status of the families as indicated by such items as nativity, education, and occupation of the household head. The broad implications of the results will be discussed further after these data shall have been presented.

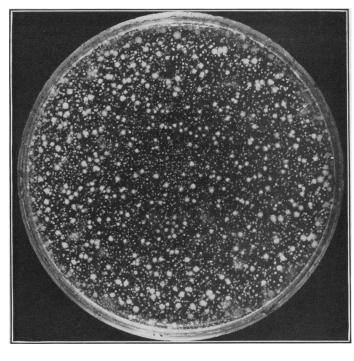
BACTERIAL CONTENT OF THE KANSAS DUST STORM ON MARCH 20, 1935

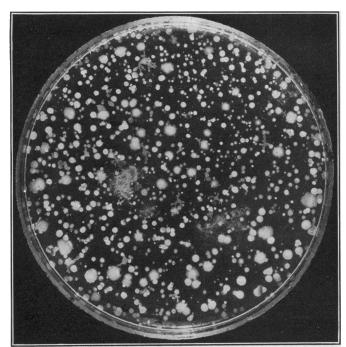
By Cassandra Ritter, Bacteriologist, Division of Sanitation, Kansas State Board of Health, Lawrence, Kans.

On March 20, 1935, there occurred a dust storm of unusual intensity, and the number of bacteria present, both outside and inside the laboratory, seemed to be a matter of such interest that they were determined by a simple experiment.

Petri dishes were prepared with sterile nutrient agar culture media. After the agar had hardened, the tops of the dishes were removed for certain lengths of time, which allowed the surface of the agar to become seeded with particles of dust. The plates were then incubated at 37° C. for 24 hours.

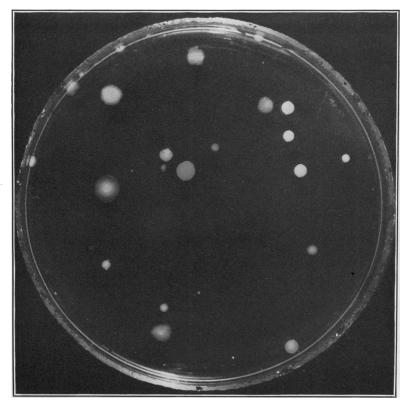
The outside exposures were made at the south entrance of Marvin Hall, University of Kansas, at Lawrence, where there was no obstruction to the wind. The exposures were made between 3 and 3:20





March 20, 1935: Exposure, 30 seconds.

March 20, 1935: Exposure 5 minutes.



March 25, 1935: Exposure, 5 minutes.

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o'clock in the afternoon, after the storm had been in progress for several hours. Exposure times were 15 and 30 seconds, and 1, 1½, 2, 3, 5, and 10 minutes. In the laboratory, plates were exposed for 20 seconds and for 1 minute, and a control plate was not exposed.

It was possible to count the colonies on only a few plates. Those with longer exposures were not only too crowded, but it was obvious that all the organisms falling on the surface did not have a chance to develop. The counts that could be made were as follows:

15 seconds, duplicate plates	600 and 650 bacteria colonies.
30 seconds	1,100 bacteria colonies.
20 seconds, inside exposure	56 bacteria colonies.
1 minute, inside exposure	95 bacteria colonies.
Control plate, inside exposure	28 bacteria colonies.

As a matter of interest, the number of bacteria falling on 1 square foot per minute was computed. Using the number 600 falling on a Petri dish of measured area in 15 seconds, we calculated 31,000 bacteria per square foot per minute.

The colonies of bacteria on the plates appeared very similar to those formed by soil organisms, some of which will appear on plates made from raw waters. This was borne out by a microscopical examination of a number of colonies. Of 11 colonies examined, all but 2 had formed spores in 24 hours; they were all rather large bacillus forms, and most of them were Gram-positive. No coccus forms were found, either in that or later microscopical examinations. This strongly indicated that the bacteria surviving in the dust were resistant soil types.

In order to show the contrast between the number of bacteria present in the air during the dust storm and the number normally present, plates were exposed in the same location and at the same time on March 25. The day at the time of exposure, 3 o'clock, was clear and calm, although dust clouds had been visible in the morning. Plates exposed 1 minute and 5 minutes showed counts of 12 and 30, respectively. A plate exposed inside for 1 minute showed a count of 12.

DEATHS DURING WEEK ENDED APRIL 13, 1935

[From the Weekly Health Index, issued by the Bureau of the Census, Department of Commerce]

	Week ended Apr. 13, 1935	Corresponding week,
Data from 86 large cities of the United States: Total deaths Deaths per 1,000 population, annual basis Deaths under 1 year of age. Deaths under 1 year of age per 1,000 estimated live births Deaths per 1,000 population, annual basis, first 15 weeks of year Data from industrial insurance companies: Policies in force. Number of death claims. Death claims per 1,000 policies in force, annual rate. Death claims per 1,000 policies, first 15 weeks of year, annual rate.	8, 438 11. 8 579 53 12. 7 67, 734, 319 13, 248 10. 2	8, 874 12. 4 675 63 12. 6 67, 698, 617 14, 298 11. 0

PREVALENCE OF DISEASE

No health department, State or local, can effectively prevent or control disease without knowledge of when, where, and under what conditions cases are occurring

UNITED STATES

CURRENT WEEKLY STATE REPORTS

These reports are preliminary, and the figures are subject to change when later returns are received by the State health officers

Reports for Weeks Ended Apr. 20, 1935, and Apr. 21, 1934

Cases of certain communicable diseases reported by telegraph by State health officers for weeks ended Apr. 20, 1935, and Apr. 21, 1934

	Diph	theria	Infl	uenza	Me	asles	Menin men	gococcus ingitis
Division and State	Week ended Apr. 20, 1935	Week ended Apr. 21, 1934	Week ended Apr. 20, 1935	Week ended Apr. 21, 1934	Week ended Apr. 20, 1935	Week ended Apr. 21, 1934	Week ended Apr. 20, 1935	Week ended Apr. 21, 1934
New England States:								
Maine		1	8		109	14	0	0
New Hampshire	l				2	167	O	1
vermont	1 2	- -			46	53	0	0
Massachusetts	3 2	14			453	1, 953	3	2
Rhode Island Connecticut		1	6	2 2	343	3	1	0
Middle Atlantic States:	, z		0.	2	1,065	52	1	. 1
New York	33	62	19	1 10	3, 156	1, 227	24	
New Jersey	12	16	15	16	1, 244	657	3	1
Pennsylvania	35	36	10	- 40	3, 044	4. 033	6	3
East North Central States:	"	•			3,011	2,000	١٠	3
Ohio	49	31	19	14	1.549	1, 207	11	4
Indiana	20	15	22	14	365	1, 073	4	ī
Illinois	29	31	46	21	3, 197	1,813	23	15
Michigan	5	17	2	i i	6, 488	251	5	2
Wisconsin	ĭl	3	6	24	1, 555	1, 595	ĭ	2
West North Central States:	_	- I		- - 	2,000	2,000	- 1	-
Minnesota	6	3	3		615	231	1	0
Iowa	8	11	3	4	537	240	4	ŏ
Missouri	44	34	103	49	776	936	8	, ă
North Dakota	5	1	13	2	31	152	Õ	Ō
South Dakota	6	3	1		68	336	0	Õ
Nebraska	5	1		10	365	232	ō l	Ŏ
Kansas		9	8	2	1, 372	510	2	Ó
South Atlantic States:	i	i	1	1		i		
Delaware	1	1			13	102	0	0
Marviena	5	9	7	8	49	1,909	6	·Õ
District of Columbia	15	7	2	2	92	226	5	2 2
Virginia	11	18			735	1,400	7	2
west virginia	17	19	37	64	317	89	1	8
North Carolina	11	16	10	17	223	2, 298	1	1
South Carolina	6	7	157	372	39	708	1	0
Georgia 2	4	6				592	1 0	1
Florida	2	9 1	2	2	81	1, 187	0	0

See footnotes at end of table.

Cases of certain communicable diseases reported by telegraph by State health officers for weeks ended Apr. 20, 1935, and Apr. 21, 1934—Continued

Joi weeks elided A	pr. 20,		**************************************		1004	Contin	ueu	
	Dipl	htheria	Infi	uenza	М	easles	Menin men	gococcus ingitis
Division and State	Week ended Apr. 20, 1935	Week ended Apr. 21, 1934	Week ended Apr. 20, 1935	Week ended Apr. 21, 1934	Week ended Apr. 20, 1935	Week ended Apr. 21, 1934	Week ended Apr. 20, 1935	Week ended Apr. 21, 1934
East South Central States: Kentucky	16 5 12	9 5 17 6	20 40 73	6 39 53	514 19 214	185 816 881	4 6 2 2	
Arkansas Louislana ³ Oklahoma ⁴ Texas ³	19	1 18 5 79	18 4 58 301	7 6 39 169	70 35 91 185	65 349 240 942	1 0 4 6	3 1 0 0 2
Mountain States: Montans	1 2	3 3 2	277 3 6	110 2	609 4 120 233 27	40 36 90 352 162	0 0 0 1	0 0 0 1
Arizona Utah ³ Pacific States: Washington Oregon California	3 1 7 30	5 42	33 62	14 5 37 36	23 10 342 205 1, 413	58 256 196 87 942	3 1 4	0 0 2 0 3
Total	497	580	1, 133	1, 161	32, 046	30, 943	154	
First 16 weeks of year	10, 985	13, 021	96, 179	40, 248	420, 741	408, 544	2, 138	903
Division and State	Week ended	Week ended Apr. 21, 1934	Week ended	Week ended Apr. 21, 1934	Week ended Apr. 20, 1935	Week ended Apr. 21, 1934	Week ended	Week ended Apr. 21, 1934
New England States: Maine New Hampshire Vermont Massachusetts Rhode Island Connecticut Middle Atlantic States:	0 0 0 0	0 0 0 0	6 9 7 237 7 110	11 12 11 225 22 91	0 0 0 0 0	0 0 0 0	1 0 0 5 0	1 0 0 3 0
New York	0 2 0	0	1, 241 173 548	874 212 741	0 0 0	0 0 0	10 0 3	8 4 11
Ohio	1 0 0 0 0	1 0 2 1 0	773 168 1, 251 352 410	796 169 610 803 242	3 0 0 0 14	0 0 5 1 50	5 2 18 2 2	5 7 4 1 2
West North Central States: Minnesota Iowa Missouri North Dakota South Dakota Nebraska Kenasa	0 0 0 0 0 0	1 0 1 0 0 1	339 81 69 66 8 57 70	66 55 95 24 4 49 39	0 18 2 0 5 33 17	7 4 7 0 6 2 11	0 0 4 0 0 1 2	1 0 8 0 1 0 2
Kansas South Atlantic States: Delaware Maryland District of Columbia Virginia West Virginia North Carolina South Carolina Georgia 2 Florida	0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	0 0 0 0 1 0	7 108 90 26 57 14 6 5	8 58 14 29 78 23 8 10 3	0 0 0 0 0 2 0	0 0 0 0 0 2 0 0	0 7 0 11 3 7 1 11 8	1 7 1 5 20 1 0 16 7

See footnotes at end of table.

Cases of certain communicable diseases reported by telegraph by State health officers for weeks ended Apr. 20, 1935, and Apr. 21, 1934—Continued

	Polion	nyelitis	Scarle	t fever	Sma	llpox	Typho	id fever
Division and State	Week ended Apr. 20, 1935	Week ended Apr. 21, 1934						
East South Central States:	_							
Kentucky	0	0	28	43	0	0	8	0 2 3 1
Tennessee	O O	1	25	26	0	1		2
Alabama	0	0	8	9	0	0	1 1	3
Mississippi 3		0	5	8	0] 1	1	1
Arkansas				۱ .			i .	_
Louisiana 1	0	0	!	3	1	1 1	.1	1
Oklahoma 4	0	0	.4	24	0	9	18	20
Texas 3	0	Q	11 50	9 81	1 11	8 36	6	.4
Mountain Ctates	וייייו	U	5 U	91	11	80	6	14
Mountain States: Montana	اما			ا ما			اما	_
Montana	0	1	5	8	5	0	0	0
Idaho s Wyoming s	ă	0	21		, 1	9	0	Ų
Colorado	ň	Ň	21 215	. 8	15	Ŏ	0	1 2
New Mexico		ŏ		31	0	0	0	
		ŏ	14	22	1		5	4
Arizona Utah 3	ואָ	ŏ	55	15	0	0	9	4 2 0
Pacific States:	١٠	U	135	11	U	6	١٠	U
	ا ما		40			_		
Washington	0	0	48	31	15	8	1	4
Orecon 5 California	0	.0	58	50	2	9	1	1
Camornia	2	10	205	213	8	2	6	6
Total	8	22	7, 193	5, 974	150	182	163	181
First 16 weeks of year	386	328	115, 048	97, 044	3, 068	2, 383	2, 103	2, 409

SUMMARY OF MONTHLY REPORTS FROM STATES

The following summary of cases reported monthly by States is published weekly and covers only those States from which reports are received during the current week.

State	Menin- gococ- cus menin- gitis	Diph- theria	Influ- enza	Malaria	Measles	Pel- lagra	Polio- mye- litis	Scarlet fever	Small- pox	Ty- phoid fever
January 1935 Colorado New Hampshire February 1935	1	34 7	17 5		2, 795		1 0	1, 049 48	10 0	5 0
Colorado New Hampshire	8	43	30		3, 457		1 0	1, 206 32	15 0	4 0
Illinois Maine Maryland Maryland Michigan Minnesota New Jersey Ohio Oregon Pennsylvania South Carolina South Carolina South Dakota Tennessee Texas West Virginia Wyorning	88 1 15 8 9 10 59 7 18 23 5 32 25 12	250 7 7 23 56 31 89 220 1 200 99 24 60 279 62 5	266 254 702 27 86 89 278 489 1, 775 20 851 5, 217 495	12 1 5 	13, 448 1, 170 389 16, 266 7, 126 5, 388 6, 471 473 22, 110 248 273 462 855 2, 189 741	2 	2 2 1 1 5 3 3 1 2 0 1 0 5 1	5, 187 79 458 1, 929 916 786 4, 735 254 2, 757 21 65 127 429 388 95	2 0 0 0 0 55 0 0 0 16 0 77	28 8 9 7 10 15 4 17 8 0 9 60 16

New York City only.
 Typhus fever, week ended Apr. 20, 1935, 6 cases, as follows: Georgia, 1; Louisiana, 1; Texas, 4.
 Week ended earlier than Saturday.
 Exclusive of Oklahoma City and Tulsa.
 Rocky Mountain spotted fever, week ended Apr. 20, 1935, 5 cases, as follows: Idaho, 2; Wyoming, 2; Oregon, 1.

January 1935	Cases	March 1985—Continue	d Cases	March 1935—Continue	d Cases
Colorado:		Food meleonines	Ca363	Cantle same threat Contd	Cases
Chicken pox	519	Food poisoning:	7	Septic sore throat—Contd.	01
Impetigo contagiosa	8 131	OhioGerman measles:	•	Maryland	21
Mumps			E 780	Michigan	88
Tetanus	1	Illinois	252	Ohio	297
Trachoma	1	Maine Maryland	135	Oregon	17
Vincent's infection		Maryland	1 576	Tennessee	14
Whooping cough	96	New Jersey	1, 0/0	Wyoming	9
		Ohio	9 400	Tetanus:	•
February 1935		Pennsylvania	5, 429	Illinois	2 2
O 3 1		Tennessee	0	New Jersey Ohio	î
Colorado:	400	South Carolina	46	Trachoma:	1
Chicken pox	482 2		20		708
Impetigo contagiosa	155	Impetigo contagiosa:	1	Illinois	765 5
Mumps Vincent's infection	133	Illinois	Å	Michigan	î
Vincent's injection	81	Orogon	31	Ohio South Dakota	5
Whooping cough	91	Oregon Tennessee	1	Tennessee	30
3. Comph. 1095		Toundies coute infectious:		Trichinosis:	30
March 1936		Jaundice, acute infectious:	6	Illinois	8
Actinomycosis:		Michigan	U	Mulno	ŝ
Pennsylvania	1	Lead poisoning:	9	Maine Maryland	,
South Dakota	ĭ	New Jersey	í	New Jersey	
Anthrax:	_		2	Ohio	1 2 7
Pennsylvania	2	Ohio Mumps:	-	Pennsylvania	2
Chicken pox:	_	Illinois	699	l	•
Illinois	2, 280	Maine	53	Tularaemia:	_
Maine	265	Maryland	123	Illinois	5
Morvland	818	Michigan	977	Maryland	3
Michigan	2. 103	New Jersey	723	Michigan	2
Minneente	440	Ohio	2 007	New Jersey	2
New Jersey	2. 152	Oregon	951	South Carolina	1
Ohio	2,774	Pennsylvania		Tennessee	5
Oregon	293	South Carolina	342	Typhus fever:	
Pennsylvania		South Dakota	238	Tennessee	1
South Carolina	93	Tennessee	197	Texas	18
South Dakota	27	Texas	568	Undulant fever:	
Tennessee	317	West Virginia	418	Illinois	7
Texas	967	Wyoming	10	Maine	5
West Virginia	204	Ophthalmia neonatorum:		Maryland	ĭ
W yoming	34	Illinois	4	Michigan	8
_		Maryland	i	Minnesota	12
Dengue: South Carolina	1	Minnesota	ī	New Jersey	
	3	New Jersey	i	Ohio	2 5
Texas		Ohio	68	Oregon	ĭ
Diarrhea and enteritis:	_	Pennsylvania	4	Pennsylvania	î
Maryland	. 2	South Carolina	14	South Carolina	ī
Ohio	12	Tennessee	2	Tennessee	ī
Ohio South Carolina	261	Paratyphoid fever:	_	Texas	2
Dysentery:		Illinois	1		•
Illinois (amoebic)	12	Maine	ī	Vincent's infection:	16
Illinois (amoebic carri-		Maryland	1	Illinois	11
ers)	32	Michigan	1	Maine	15
Illinois (bacillary)	3	Oregon	2	Maryland	23
Maryland (bacillary)	1	Texas	3	Michigan	8
Michigan (amoebic)	2	Puerperal septicemia:		Oregon	8
Minnesota (amoebic) Minnesota (bacillary)	4	Illinois	5	Tennessee	
Minnesota (bacillary)	3	Ohio	9	Whooping cough:	
Ohio	3	Rabies in animals:	1	Illinois	1,075
Pennsylvania	1	Illinois	37	Maine	141
Tennessee	2	Maryland	6	Maryland	199
Texas	15	New Jersey	7	Michigan	1,0/3
Epidemic encephalitis:		Oregon	2	Minnesota	163
Illinois	10	Oregon South Carolina	73	New Jersey	1,0/2
Michigan	ĭ	Rocky Mountain spotted		Ohio	755
Minnesota	4	fever:		Oregon	121
New Jersey	4	Oregon	2	Pennsylvania	1, 4/8
Ohio	7	Scables:	_ [South Carolina	152 39
Oregon	7 2 7	Maryland	2	South Dakota	230
Pennsylvania	7	Oregon	44	Tennessee	483
South Carolina	3	Septic sore throat:	٠	Texas West Virginia	207
Tennessee	1	Illinois	19	west virginia	49
Texas	1	Maine	1	Wyoming	20

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WEEKLY REPORTS FROM CITIES

City reports for week ended Apr. 13, 1935

[This table summarizes the reports received regularly from a selected list of 121 cities for the purpose of showing a cross section of the current urban incidence of the communicable diseases listed in the table. Weekly reports are received from about 700 cities, from which the data are tabulated and filed for reference]

State on 3 site.	Diph-	٠,	Influenza		Pneu- monia	Scar- let	Small-	Tuber-	Ty- phoid	Whoop- ing	Deaths,
State and city	cases	Cases	Deaths	sles cases	deaths	fever cases	cases	deaths	fever cases	cough cases	causes
Maine: Portland	0		0	0	6	6	0	0	2	0	29
New Hampshire:	0			0	1 1		1			l	
Concord Nashua Vermont:	ŏ		0	0	2	3	0	0	0	3 0	13
Barre Burlington Massachusetts:	0		0	69	0	1	0	ō	0	0	15
Boston	2 1		0	34 16	25 7	54 3	0	7	1 0	13 2	214
Fall River Springfield	0		0	134	1 1	11	0	1 1	0	10	30 47
Worcester Rhode Island:	0		0	5	10	22	0	1	0	5	62
Pawtucket	1		ļ	.0	0	0	0	o l	0	.0	14
Providence Connecticut:	0		0	141	7	9	0	3	0	16	75
Bridgeport Hartford	0	1	0	28	10	9 16	0	1 1	0	3 18	32 60
New Haven	ō		ŏ	634	4	ĭ	ŏ	ō	ŏ	ő	46
New York:											
Buffalo New York	1 25	5	1 3	153 1, 472	13 153	53 848	0	7 92	0 5	20 223	140 1, 465
Rochester	0		0	245	2	14	0	1	0	30	64
Syracuse New Jersey:	0		0	431	5	8	0	1	0	26	60
Camden Newark	3	2 6	0	489	5 11	5 11	0	0 7	0	0 80	38 116
Trenton	ŏ		ŏ	23	5	. 5	ŏ	5	ŏ	2	49
Pennsylvania: Philadelphia	8	11	7	36	51	121	0	23	اه	78	499
Pittsburgh	8	8	4	507	25	46	0	11	Ö	21	182
Reading Scranton	1		0	62 56	2	9	0	2	0	1 0	23
Ohio:										-	
Cincinnati Cleveland	4 9	53	2 2	500	14 15	31 52	0	9 14	0	0 33	122 194
Columbus	2		0	166	6	36	0	3	0	5	69
ToledoIndiana:	0	2	1	98	5	14	0	4	1	19	75
Fort Wayne	3		1	14	3	2	o l	1	0	1	. 23
Indianapolis South Bend	0		0	77 3	20 3	20 8	0	8	0	22	115 20
Terre Haute Illinois:	1		0	0	0	1	0	0	0	0	23
Chicago	19	5	3	1, 568	57	675	0	36	2	61	724
Springfield Michigan:	0		1	22	2	19	0	1	0	12	27.
Detroit Flint	1 0	2	2 1	2, 832 51	24	145	0	15	1 0	123	267 27
Grand Rapids.	ŏ		î ¦	140	4	10	ŏ¦	ō	ŏ ¦	31	43
Wisconsin: Kenosha	0		o	73	1	34	0	0	0	5	6
Milwaukee	Ö	1	• 1	141	7 1	139	Ō	5	0	42	107
Racine Superior	0		0	71 80	1	14	0	0	0	7	9 7
Minnesota:										-	
Duluth.	0		0	437	6	0 166	0	0	0	1	26 97
Minneapolis St. Paul	3 3	i	0	489 13	10 7	43	0	0	0	27 13	97 66
lowa: Davenport	0	1		1	- 1	1	0		0	0	
Des Moines	2 2			396		5	0		0	0	34
Sioux City Waterloo	3	2		3 2		1 2	0		0	10	
Missouri: Kansas City	8	_ [,	130	10	7	0	5	0		103
St. Joseph	0		0	5	1	0	0	1	2	2 2 3	8
St. Louis	12 i		2	24	15	12	0 1	13	1	3	205

City reports for week ended Apr. 13, 1935—Continued

	Dinh Influenza						T		Tru.	Whoop-	
State and city	Diph- theria	100	luenza	Mea- sles	Pneu- monia	Scar- let fever	Small- pox	culosis	Ty- phoid fever	ing cough	all
	cases	Cases	Deaths	cases	deaths	cases	cases	deaths	cases	cases	causes
North Dakota:				1 _							
Fargo Grand Forks	1 0		1	7	0	10	0	0	0	0 2	
South Dakota:	1			l		ļ	1		İ	1	
Aberdeen	0			23		0	0		0	0	
Nebraska: Omaha	2	l	1	74	9	10	1	0	0	0	57
Kansas:		i		ļ	l]	Ì				l
Topeka Wichita	0		0	490	2	5	ō	Ö	0	1	32
Delaware:	1	1		ł	1	I					i
Wilmington	0		0	4	6	9	0	0	0	1	22
Maryland: Baltimore	1	4	2	32	26	51	0	14	1	22	201
Cumberland	l o	ļ	0	8	0	1	Ó	0	0	0	9
Frederick District of Col.:	0		0	0	0	0	0	0	0	0	4
Washington	16	2	1	50	18	74	0	21	0	4	165
Virginia:	0	1	0	17	١.	1	0	1	0	30	٠.,
Lynchburg Norfolk	l ö	i	ľ	41	1 5	2	8	Ì	l ŏ	50	16 44
Richmond	Ó		2	129	4	0	0	1	0	0	57
Roanoke West Virginia:	1		0	15	1	0	0	0	0	0	19
Charleston	0		0	8	3	2	0	2	0	0	23
Huntington	1 0			3		3	0		l o	0 3	
Wheeling North Carolina:	ľ		1	60	4	•	0	۰	0	1 3	16
Raleigh	0		0	0	2	Ō	0	1	0	0	5
Wilmington Winston-Salem	0		8	0 3	2 2	1 2	0	0	0	6	11
South Carolina:	ľ		"	1			١ ،				11
Charleston	Ŏ	10	<u>-</u> -	3	2	0	0	2	0	0	22
Columbia Greenville	0		0	0	1 1	0	0	0	0	8	19 16
Georgia:			1								1
Atlanta	4	23	0	0	8	2	0	4	0	0	62
Brunswick Savannah	0	2	0 2	0	0	1	0	i	0	i	5 26
Florida:			1								
Miami Tampa	1	2	0	68	1 1	0 2	0	1 1	0	4 3	23 27
Kentucky:			l								
Ashland	0		0	15	1	0	0	0	0	0	
Lexington Louisville	1 1		8	20 442	11	0 19	0	0 2	0	47	19 66
Tennessee:		•	i	112			l .				
Memphis	0		0	1	7	.4	0	3	1	5	69
Nashville Alabama:	0		1	1	6	14	0	1	0	2	60
Birmingham	3	6		20	6	1	0	4	0	9	57
Mobile	0		1	1 27	2	1	0	3	0	0	29
Montgomery		[١	١			7	
Arkansas:	_			ا		ا م	0				
Fort Smith Little Rock	0 1		0	0 26	6	0	ŏ	·ō	0	1	7
Louisiana:				1							
New Orleans	17	2	2	49	8 7	7	0	9 5	7	1 0	131
Shreveport Texas:	1		U	3	'	ľ		١	٠	۰	51
Dallas	5	2	1		4	4	0	1	0	1	42
Fort Worth	0		0		3 3	0	8	3	2 0	0	35 11
Houston	4		ŏ	4	4	ο	ĭ	4	ŏ	ŏ	64
San Antonio	0		2	0	4	1	0	9	0	1	66
Montana:						1	I	l	1	ļ	
Billings	0		0	14	0	0	0	0	0	0	7
Great Fails			0	18				0		i-	5
Helena Missoula	ŏ		ŏ	150	ŏ	ĭ	ŏ	ŏ	ŏ	ō	5
Idaho:			١		.	ا ۽	ا ا	ا	ا ۾	ا	
Boise	0		0	1	1	3	0	0	0	0	8
Denver	4	43	0	147	4	136	2	6	<u>-</u> -	.6	93
Pueblo	0 1	'	0	116	0 1	3 1	0 1	1 1	0 1	15	6

Indianapolis.

Springfield

Illinois: Chicago

Detroit. Wisconsin:

Milwaukee_

Minneapolis ...

Iowa:
Davenport
Sioux City

Michigan:

Minnesota:

City reports for week ended Apr. 13, 1935-Continued

State and city	Diph- theris		Deaths	Mea- sles cases	Pneu- monia dea th	Scar- let fever	Small- pox cases	Tuber- culosis deaths	fever	Whoop- ing cough cases	Deaths, all cases
New Mexico: Albuquerque. Utah: Salt Lake City. Nevada: Reno	0 0 0 0 0 0 0		0 1 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	1 7 2 130 173 3 121 0 65 113	1 5 0 1 1 3 8 6 14 1	0 84 0 13 7 4 9 1	0 0 0 3 3 0 0	2 1 0 3 1 0 0	0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	1 115 0 7 0 5 0 0 0 11 0 0	10 83 4 84 45 36 75
San Francisco	2	Maning	gococcus	24	ıi	18	0	14	i	16 cococcus	196
State and city			Deaths	Polio- mye- litis cases	State and city			7		ngitis	Polio- mye- litis cases
Rhode Island: Providence Connecticut: Hartford		0	1		Nebraska:		ph		2	0	0
New York: New York Rochester New Jersey: Newark		6 2 0	9 0	(Mai 1 Dis	Maryland: Baltimore District of Columbia: Washington			4	0	1
Pennsylvania: Pittsburgh Ohio: Cincinnati Cleveland		1 8 2	2 3 0		0 Ker	Norfoli tucky: Louisv	: ille lle		1 2 1	1 0 0	0 0 0
Toledo Indiana: Indianapolis		2 1	ž 0	(0 Ala	bama:	gh am	1	1	0	0

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Louisiana:
New Orleans
Washington:
Seattle

Oregon:
Portland -California:

Spokane....

Los Angeles Sacramento

San Francisco.....

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Epidemic encephalitis.—Cases: New York, 16; Cleveland, 2; Toledo, 1; St. Paul, 1. Pellagra.—Cases: Winston-Salem, 2; Charleston, S. C., 3; Atlanta, 1; Tampa, 1. Typhus fever.—Cases: New York, 1; Atlanta, 1.

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FOREIGN AND INSULAR

CEYLON

Malaria.—A report dated March 1, 1935, states that the peak of the malaria epidemic was thought to have been passed in Ceylon. A severe drought in many parts of the island was causing additional anxiety. The following mortality figures were given, showing the great increase in deaths (all causes) which occurred during the epidemic.

Numb of deat	ths	Nun of dec	aths
November 1933 9, 4 December 1933 9, 0 January 1934 11, 5	149	December 1934	738

CUBA

Provinces—Notifiable diseases—4 weeks ended April 6, 1935.—During the 4 weeks ended April 6, 1935, cases of certain notifiable diseases were reported in the Provinces of Cuba, as follows:

Disease	Pinar del Rio	Habana	Matan- zas	Santa Clara	Cama- guey	Oriente	Total
Cancer Chicken pox Diphtheria	1	2 5	4	4 1 1	4	1 1	7 11 6
Hookworm disease Leprosy. Malaria. Measles	188	13	42 7	778 778 34	137	21 460 2	23 1, 605 56
Poliomyelitis Tuberculosis Typhoid fever	1 4	5 1	24 8	70 21	12 20	51 6	3 166 56

CZECHOSLOVAKIA

Communicable diseases—February 1935.—During the month of February 1935, certain communicable diseases were reported in Czechoslovakia, as follows:

Disease	Cases	Deaths	Disease	Cases	Deaths
Anthrax Cerebrospinal meningitis Chicken pox Diphtheria Dysentery Infinenza Lethargic encephalitis Malaria	6 25 275 2,554 11 41,747 2 6	1 6 1 195 4 37 1	Paratyphoid fever	5 8 42 1,630 81 309 18	2 15 20 33

ITALY

Communicable diseases—4 weeks ended December 9, 1934.—During the 4 weeks ended December 9, 1934, certain communicable diseases were reported in Italy, as follows:

	Nov. 12-18		Nov. 19-25		Nov. 26-Dec. 2		Dec. 3-9	
Disease	Cases	Com- munes affected	Cases	Com- munes affected	Cases	Com- munes affected	Cases	Com- munes affected
Anthrax Cerebrospinal meningitis Chicken pox Diphtheria and croup Dysentery Lethargic encephalitis Measles Poliom yelitis Scarlet fever Typhoid fever	21 13 263 658 11 6 1, 382 13 511 604	20 13 113 377 10 6 256 10 221 359	22 10 417 872 8 3 1,808 14 550 655	21 9 130 380 6 3 252 13 203 351	12 12 432 898 9 1 1,857 16 515 559	12 11 144 440 8 1 300 14 185 324	15 13 345 826 10 1 2,000 7 476 563	14 12 118 384 7 1 292 7 190 319

YUGOSLAVIA

Communicable diseases—March 1935.—During the month of March 1935, certain communicable diseases were reported in Yugoslavia, as follows:

Disease	Cases	Deaths	Disease	Cases	Deaths
Anthrax Cerebrospinal meningitis Diphtheria and croup Dysentery Erysipelas Influenza Measles	29 17 559 16 155 70, 620 1, 787	2 6 60 1 7 109 34	Paratyphoid fever	5 192 13 16 159 117	2 7 10 20 7

CHOLERA, PLAGUE, SMALLPOX, TYPHUS FEVER, AND YELLOW FEVER

(NOTE.—A table giving current information of the world prevalence of quarantinable diseases appeared in the Public Health Reports for Apr. 26, 1935, pp. 580-594. A similar cumulative table will appear in the Public Health Reports to be issued May 31, 1935, and thereafter, at least for the time being, in the issue published on the last Friday of each month.)

Plague

British East Africa—Kenya.—During the week ended March 16, 1935, 1 case of plague was reported at Kenya, British East Africa.

Indo-China—Island of Nao-Tchao.—During the period March 1-10, 1935, 20 cases of plague with 15 deaths were reported in the Island of Nao-Tchao, Indo-China.

Yellow Fever

Sierra Leone—Freetown.—On March 10, 1935, 1 case of yellow fever was reported at Freetown, Sierra Leone.